From Population Growth to TFP Growth

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From Population Growth to TFP Growth*

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Abstract
A slowdown in population growth causes a decline in business dynamism by increasing the share of old businesses. But how does it affect productivity growth? We answer this question by extending a standard business dynamics model to include endogenous productivity growth. Theoretically, the growth rate of the size of surviving old businesses is a "sufficient statistic" for determining the direction and magnitude of the impact of population growth on productivity growth. Quantitatively, this effect is significant across balanced growth paths for the United States and Japan. TFP growth in the United States falls by 0.3 percentage points because of the slowing in population growth between 1970 and 2060. The same driving force produces a significantly bigger response in Japan. Despite the significant long-run effect, we discover that changes in TFP growth are slow in reaction to population growth changes due to two short-run counterbalancing factors.

Keywords: population growth, economic growth, firms dynamics, demographics, productivity, innovation, TFP, Japan.

JEL Codes: E20, J11, O33, O41.

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1 Introduction

In developed countries, there is an increasing concern that slowing population growth may lead to a decline in economic growth.\footnote{Although we usually refer to it as population growth, the driving force in our study is labor force growth.} Figure 1 shows the estimated trends for the United States and Japan, the countries that we will consider in our quantitative exercises. At the beginning of our data in 1900, or in the 1970s when the baby boomers entered the job market, the trend in US labor force growth in the US was close to 2.2%. In contrast, current forecasts estimate that in 2060, it will be less than 0.3%. The fall is more dramatic in Japan, and it started earlier. While trend labor force growth was close to 1.5% in 1950, it is expected to be below -1.3% in the 2040s.

Figure 1: Labor force growth rate

![United States and Japan labor force growth](source)

Several recent studies (Karahan, Pugsley and Sahin, 2019; Peters and Walsh, 2021; Hopenhayn, Neira and Singhania, 2022) demonstrated that a slowing in population growth in the United States led to a decline in business dynamism by increasing the share of old businesses. But how does this shift in population and business demographics affect productivity growth? We aim to provide an answer...

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to this specific question.

This paper incorporates population growth and endogenous productivity growth into a business dynamics model to achieve this goal. As in Hopenhayn (1992), businesses begin with low productivity and increase their productivity throughout their life cycle. However, the productivity of younger businesses is determined by the intensity of innovation within the business, which is endogenous. A business innovation improves on previous innovations, as in Romer (1990) and Aghion and Howitt (1992). This innovation is the first driver of growth in our model. Older businesses’ productivity also increases with age. This productivity growth represents the second growth engine in our model and captures technology advancements made by mature, successful businesses. Therefore, for our model to display balanced growth, one of these two forces—innovation by young businesses or technological advancements by established and successful businesses—must be present. In addition, every period, some incumbent businesses exit, and new businesses enter.\(^2\)

The main theoretical result is that in comparing BGPs, the shape of the businesses’ life-cycle profile determines the sign and magnitude of the impact of population growth on productivity growth. In particular, we identify a “sufficient statistic”—namely, the growth rate of surviving old businesses. If the growth rate of the size of surviving old businesses is negative, a fall in the population growth rate will result in a decrease in the rate of aggregate productivity growth. Two components make up the mechanism for this result. First, as population growth declines, so does the growth rate in the number of businesses; otherwise, the average firm size will diverge. As a result, an economy with a lower population growth rate will have a lower proportion of young businesses as a small number of new businesses relative to existing businesses implies that most businesses are old. The second element of the mechanism underlying the aforementioned result is

\(^2\)We assume exit is exogenous for the theoretical characterization of the model, but we relax this assumption in a subsequent section.
based on the productivity growth of old businesses relative to overall productivity growth. The “sufficient statistic” is precisely a measure of these two growth rates. If the size of surviving old businesses decreases over the life cycle, their productivity is expanding at a slower rate than the average productivity of the economy. Thus, putting these two factors together, as the labor force expands more slowly and the proportion of old businesses grows, if the productivity growth of old businesses is lower than the average, the total productivity growth will be lower.

These theoretical results are feasible because there is a fully elastic supply of startups, entrants’ innovation is constant, and we assume that survival rates and the growth rate of productivity of successful businesses are exogenous. Subsequently, we examine computer-solved versions of the model that relax these assumptions. The main quantitative findings are based on comparing BGP’s with different population growth rates in the US and Japan. For these exercises, we consider not only the model analyzed in the theoretical section of the paper. Our benchmark model for all the quantitative exercises is a model extended to incorporate congestion at entry and spillovers from new to old businesses’ productivity growth. However, we also analyze the impact of extending the model to incorporate endogenous exit and endogenous innovation by mature businesses on the BGP analysis. We conclude that population growth has a significant impact on productivity growth. In the benchmark case, a drop in population growth, as projected for the US for 1970-2060, implies a long-run decline in productivity growth of about 0.3 percentage points. Similarly, for Japan, the predicted drop in population growth for 1950-2060 implies, in the long run, a 0.6 percentage point reduction in productivity growth.

Next, we compute transitional dynamics for the economies calibrated to the US and Japan. Since computing a transition is computationally challenging, we focus primarily on the benchmark model extended to include congestion at entry and spillovers from new to old businesses’ productivity growth. The main experiment
involves giving the model time series for the trend in labor force growth and sim-
ulating the evolution of total productivity (TFP) growth. An essential result of the
impact of population growth on TFP growth is that it takes a long time to occur.
We also investigate why the response of TFP growth is sluggish and discover two
significant factors: a labor-reallocation effect and a level-vs-growth effect. Both ef-
effects fade in the long term, leading to our main result comparing BGPs; however,
in the meantime, they partially offset the drop in TFP growth.

The paper’s final section validates the mechanism proposed here. The dynamic
correlation between labor force growth and productivity growth produced by the
model is very similar to the correlation found in data for US states. This result,
obtained using local projections, supports the proposed mechanism and its quan-
titative significance. Furthermore, instrument variable regressions suggest a causal
effect of labor force growth on productivity growth.3

A recent paper, Alon, Berger, Dent and Pugsley (2018), is related to our re-
search since it investigates how declining business entry and aging incumbent
businesses—the exact mechanism studied here—affect aggregate productivity growth.
They employ a highly pertinent dataset, the Census Bureau’s Revenue-enhanced
Longitudinal Business Database, to characterize business growth throughout its
life cycle. They have two significant findings in connection to our research. First,
they show that productivity growth is downward-sloping and convex throughout
the business’ life cycle. Second, they found that these profiles stay unchanged over
time. Using these findings, they compute counterfactuals and show that declining
firm entry and aging established businesses had a cumulative drag on aggregate
productivity of 3.1% since 1980.

There are two main differences between their research and our work. First,
we use establishment-level data to calibrate our model because establishments are
more connected to the idea of innovation we intend to capture. However, their

3Additionally, in section 7.3, we show that the model generates the recent slowdown in busi-
ness dynamism in the United States, which has been examined in several recent papers.
findings for firm-level data are reassuring that this choice does not drive our results and make our and their findings complementary. Second, our analysis uses an equilibrium structural model that allows for endogenous forces (innovation, entry, exit) and prices (wages, interest rates) to change as we perform counterfactuals (balanced growth paths and transitional dynamics). This choice makes our work more related to two recent papers studying the relationship between population growth and business dynamism using similar firm dynamics models. According to Karahan, Pugsley and Sahin (2019) and Hopenhayn, Neira and Singhania (2022), the slowing of labor force growth has resulted in a startup deficit, which can explain what is widely known as a reduction in business dynamism. These papers share several characteristics of our framework. As a validation exercise, we show that our model can also reproduce the reduction in business dynamism in the US. However, we focus on the impact of the same driving force on TFP growth rather than on business dynamism.

Peters and Walsh (2021) also focuses on the relationship between population growth and business dynamism, but it is more associated with our work because it also displays endogenous growth. While we purposely abstract away from scale effects to focus on our new mechanism, they use a semi-endogenous growth model. Their research supplements ours by exploring the role of product variety. They don’t investigate the mechanism discussed here because they assume a business’s innovation is independent of its age. We bring an original and quantifiable mechanism relating population growth to TFP growth that complements their work by focusing on business growth over the lifecycle. In short, Peters and Walsh (2021) argues that a slowdown in population growth reduces productivity growth by slowing down the growth of product variety, while our contention is

5Although it is not the focus of Engbom (2018)’s analysis, the transitional dynamics shown in that paper’s figure 10 reveal a slight decline in growth between 1970 and 2050.
6For a survey on the importance of scale effects, see Jones (2022). In Section 8, we argue that the magnitude of our results is smaller but comparable to the scale effects described there.
that productivity growth slows down even when there is only one product in the economy. Importantly, our model’s mechanism does not conflict with Peters and Walsh (2021), so we can integrate these two mechanisms into a unified model, resulting in the combined effect on productivity growth.\footnote{Several other papers have been written about the importance of population growth to economic prosperity. For example, Cooley, Henriksen and Nusbaum (2019) investigate this demographic change’s impact on output growth via capital accumulation and labor productivity. Vandenbergrouck (2021) investigates the slowdown in output-per-worker growth in the 1960s and 1970s.}

There are two recent studies linking population growth to productivity growth. First, Jones (2020) investigates the extreme case of long-run negative population growth in the context of models of ideas, which include a variety of endogenous and semi-endogenous growth models. He discovers that negative population growth leads to stagnant living standards as the population vanishes. Recently, Kalyani (2022) finds a negative association between inventors’ creativity and age and argues that a larger proportion of older workers in the labor force will result in lower productivity growth because inventors are, on average, less creative.

\section{Model}

The economy is made up of businesses and households. Households own businesses and make decisions about consumption and investment. Businesses are the most important part of the framework since they innovate, hire workers and rent capital. In equilibrium, slower labor-force growth reduces the number of startups. This will change business demographics, which is critical for determining the relationship between population growth and productivity growth.
2.1 Household

A representative household populates the economy and solves

$$\max_{\{c_t, k_t\}} \sum_{t=1}^{\infty} \frac{(\beta g_{M_t})t-1 c_t^{1-\epsilon}}{1-\epsilon}$$

subject to

$$c_t + g_{M,t+1} k_{t+1} = w_t + s_t - e_t + r_t k_t + (1 - \delta) k_t,$$

where $k_t \equiv K_t / M_t$ is capital per person, $s_t \equiv \sum_i S_{i,t} / M_t$ is business surplus per person, $e_t \equiv E_t / M_t$ is the initial cost of starting businesses per person, $\delta$ is depreciation rate, $\beta$ is discount factor, and $g_{M,t+1}$ is population growth rate. Note that this representative household also owns the businesses.

2.2 Businesses

There are $N_t$ businesses in the economy. They have decreasing returns to scale and solve

$$S_{i,t}(x_{i,t}, w_t, r_t) = \max_{l_{i,t}, k_{i,t}} \{x_{i,t}^{\alpha} k_{i,t}^{\beta} (1-\alpha-\beta) - w_t l_{i,t} - r_t k_{i,t}\},$$

taking as given wages $w_t$ and capital rental rate $r_t$. The solutions for $l_{i,t}$, $k_{i,t}$, and $y_{i,t}$ are linear in productivity,

$$l_i = x_i \left[ \frac{\alpha}{r} \left( 1 - \alpha - \frac{\xi}{w} \right) \right]^{\frac{1}{\beta}}, \quad k_i = x_i \left[ \frac{\alpha}{r} \left( 1 - \alpha - \frac{\xi}{w} \right) \right]^{\frac{1}{\beta}},$$

$$y_i = x_i \left[ \frac{\alpha}{r} \left( 1 - \alpha - \frac{\xi}{w} \right) \right]^{\frac{1}{\beta}}.$$

Also, the average productivity in the economy is $X \equiv \frac{1}{N} \sum_i x_i$. When average productivity is combined with the expressions above, we get useful expressions for aggregate variables that we will use later to define the economy’s equilibrium.
Therefore, output, labor, and capital can be written as

\[
\begin{align*}
L & = \left\lfloor \left( \frac{\alpha}{r} \right)^{\frac{\alpha}{1 - \alpha - \zeta}} \left( \frac{1 - \alpha - \zeta}{w} \right)^{1 - \alpha - \zeta} \right\rfloor \frac{1}{\zeta} \text{NX}, \\
K & = \left\lfloor \left( \frac{\alpha}{r} \right)^{\frac{\alpha + \zeta}{1 - \alpha - \zeta}} \left( \frac{1 - \alpha - \zeta}{w} \right)^{1 - \alpha - \zeta} \right\rfloor \frac{1}{\zeta} \text{NX}, \\
Y & = \left\lfloor \left( \frac{\alpha}{r} \right)^{\frac{\alpha}{1 - \alpha - \zeta}} \left( \frac{1 - \alpha - \zeta}{w} \right)^{1 - \alpha - \zeta} \right\rfloor \frac{1}{\zeta} \text{NX}.
\end{align*}
\]

2.3 Innovation

Innovators use the ideas of successful businesses to generate their own new ideas at any given time. A new technology takes one period to start production. Let \( \chi \) be the average productivity of successful businesses, which will be precisely defined later. An innovator will then choose an innovation step size, \( g \), which measures the difference between the innovator’s potential productivity, \( \hat{x} \), and the reference productivity, \( \chi \). Thus, the cost of research for generating \( \hat{x} \) is proportional to how far ahead of the pack the project is, \( R(\hat{x}/\chi) = \frac{1}{z_R} \left( \frac{\hat{x}}{\chi} \right)^\iota, \) with \( \iota > 2 \).

Although we assume that this cost is paid entirely at the innovation stage, one period before the firm’s birth, the growth caused by R&D realizes stochastically over time. Since there are no financing frictions, this cost could be the expected discounted investment costs paid over time with no significant change in the model.\(^8\)

After the research stage, innovators develop ideas to start their businesses. The probability of entering the market (\( \sigma \)) hinges on the amount of money spent on developing the project, \( D(\sigma) = \sigma^2/(2z_D) \).\(^9\)

The value of a project started with potential productivity \( \hat{x} \) is

\[
I(\hat{x}; \{w_t\}, \{r_t\}) = \sum_{t=1}^\infty \hat{\beta}_t \mathbb{E}_{\hat{x}}[S(x_t; w_t, r_t) | \hat{x}],
\]

\(^8\)Similarly, since there are no financing frictions, we could alternately interpret that a share of unsuccessful businesses that exit at each age \( a \) are unsuccessful businesses that choose again R&D at each age \( a \) as they would choose the same innovation step size as new businesses.

\(^9\)The option choice of \( \sigma \) by innovators, which resembles Greenwood, Han and Sanchez (2022), is unimportant for our results, but it simplifies some of the expressions.
where $\hat{\beta}_t$ is the market discount factor. At the time of innovation, an innovator chooses $\sigma$ and $\hat{x}$ to maximize its payoff,

$$V(\{w_t\}, \{r_t\}, \chi_t) = \max_{\sigma_t, \hat{x}_t} \sigma_t I(\hat{x}_t; \{w_t\}, \{r_t\}) - w_t R(\hat{x}_t/\chi_t) - w_t D(\sigma_t).$$

(3)

In partial equilibrium, solving this problem yields the innovator’s chosen step size of innovation, $g^*$, the probability of starting the project, $\sigma^*$, and the maximized payoff, $V^*$. The value $V^*$ is important because the household is willing to start a business if this value covers the initial fixed cost. As a result, in equilibrium with entry, the following free-entry condition must be met:

$$V_t \leq w_t c_E.$$  (4)

The assumption that the entry cost increases one-to-one with wages makes the model tractable and is common in growth models (e.g. Klette and Kortum, 2004). The assumption is also supported by the data presented in Klenow and Li (2022).

2.4 Life-cycle profile of productivity

This section slightly simplifies the productivity process that we use in the quantitative section to replicate various empirical facts on business dynamics.

Let us consider a project with $g\chi$ potential productivity. If the project succeeds at age=1, its productivity will equal $g\chi$; if it is unsuccessful, it will equal $\theta g\chi$, where $\theta < 1$. Each period, a fraction $\lambda$ of unsuccessful projects succeed, and their productivity increases from $\theta g\chi$ to $g\chi$. While unsuccessful, the project’s productivity remains constant. On the other hand, the productivity of successful initiatives grows at a constant pace $g_S$. The survival rates $s_S$ and $s_U$ for successful and unsuccessful businesses are different. Specifically, we assume that $s_S > s_U$ represents the larger exit rate of unsuccessful businesses compared to successful businesses. Although we do not impose this condition in the calibration section, we find that successful businesses are more likely to survive than unsuccessful businesses to

\[ \hat{\beta}_t = \prod_{j=1}^{t} \frac{1}{(1+r_j-\sigma_j)}. \]
capture the growth over the life-cycle of average business size and size of surviving businesses. Figure 2 depicts a three-period example of the business lifecycle.

Figure 2: A business’s life-cycle (example up to age 3)

The reason for this simplified structure for productivity is that it allows us to construct some useful expressions for a business life-cycle. The businesses born \( a \) years ago (i.e., those age \( a \) today) can be divided into businesses that today are (i) out of business, (ii) unsuccessful, and (iii) successful. The next two expressions represent the contribution of the last two groups to the average productivity of age \( a \) businesses relative to their potential productivity. Of course, the contribution of businesses that are currently out of business is zero.
For unsuccessful businesses, the expression is simply
\[ \Lambda_{U,a} \equiv \theta(1-\lambda)^{a}(s_{U})^{a-1}, \]
where the first part, \( \theta \), is included because unsuccessful businesses’ productivity is \( \theta \) of the potential productivity, and this term is written relative to potential productivity. The second term is the probability that businesses remain unsuccessful until age \( a \), and the last term is the probability that the business does not exit before age \( a \).

The same expression for successful businesses is more involved,
\[ \Lambda_{S,a}(g_{S}) \equiv \sum_{j=1}^{a} \left[ g_{S}^{j-1}(s_{U})^{a-j}(1-\lambda)^{a-j}\lambda(s_{S})^{j-1} \right]. \]
But also, in this case, the first part, \( g_{S}^{j-1} \), is in place to adjust the productivity relative to potential productivity. Thus, for businesses that became successful \( j \) years ago, the adjustment factor is \( g_{S}^{j-1} \), to account for the productivity growth rate since they became successful. The next four terms together make the probability that a business survived and remained unsuccessful until age \( a - j \), it became successful at age \( a - j \) and survived as a successful business from age \( a - j \) until \( a \).

Why is this notation useful? We can use it to compute the average productivity based on potential productivity, \( \hat{x}_{t} \), and the number of new businesses, \( n_{t} \). In particular, for age 1 businesses, average productivity is simply
\[ X_{1,t} = \frac{\text{sum prod of age-1 businesses}}{\text{number of age-1 businesses}} = \frac{\hat{x}_{t}(\Lambda_{S,1}(g_{S}) + \Lambda_{U,1}) n_{t}}{(\Lambda_{S,1}(1) + \Lambda_{U,1}/\theta) n_{t}} = \hat{x}_{t}(\lambda + (1-\lambda)\theta). \]
Similarly, for age 2 businesses, average productivity is simply
\[ X_{2,t} = \frac{\hat{x}_{t-1}(\Lambda_{S,2}(g_{S}) + \Lambda_{U,2}) n_{t-1}}{(\Lambda_{S,2}(1) + \Lambda_{U,2}/\theta) n_{t-1}} \]
\[ = \frac{\hat{x}_{t-1} (s_{U,1}(1-\lambda)\lambda + g_{S} s_{S,1} \lambda + \theta s_{U,1}(1-\lambda)(1-\lambda))}{(s_{U,1}(1-\lambda)\lambda + s_{S,1} \lambda + s_{U,1}(1-\lambda)(1-\lambda))}. \]
Note that we can also compute the average productivity of the pool of age-1 and
age-2 businesses as

\[ X_{1-2,t} = \frac{\text{sum prod of age-1 and age-2 businesses}}{\text{number of age-1 and age-2 businesses}} = \frac{\hat{x}_t (Λ_{S,1}(g_S) + Λ_{U,1}) n_t + \hat{x}_{t-1} (Λ_{S,2}(g_S) + Λ_{U,2}) n_{t-1}}{(Λ_{S,1}(1) + Λ_{U,1}/θ) n_t + (Λ_{S,2}(1) + Λ_{U,2}/θ) n_{t-1}}. \]

Following this logic, the average productivity of all businesses in the economy is

\[ X_t = \frac{\sum_{a=1}^{\infty} \hat{x}_{t-a+1} (Λ_{S,a}(g_S) + Λ_{U,a}) n_{t-a+1}}{\sum_{a=1}^{\infty} (Λ_{S,a}(1) + Λ_{U,a}/θ) n_{t-a+1}}. \] (5)

This equation is crucial to solving the model since it depends on two key equilibrium variables: potential productivity \( \hat{x}_t \) and number of entrants \( n_t \).

Similarly, the expected productivity at age \( a \) of entrants whose potential productivity is \( \hat{x} \) is

\[ \mathbb{E}[x_a | \hat{x}] = (Λ_{S,a}(g_S) + Λ_{U,a})\hat{x}, \] (6)

and the survival probability up to age \( a \) is

\[ Λ_{S,a}(1) + Λ_{U,a}/θ. \] (7)

### 2.5 Law of motion for the number of projects

We can write the law of motion for the number of projects (by type and total) given the number of entrants \( n_t \) at a given time \( t \) as

\[ N_{U,t} = \sum_{a} n_{t-a} Λ_{U,a}/θ, \] (8)

\[ N_{S,t} = \sum_{a} n_{t-a} Λ_{S,a}(1). \] (9)

\[ N_t = N_{U,t} + N_{S,t}. \] (10)

### 2.6 Market-clearing conditions

To close the model, three market-clearing conditions must be met. The labor-market-clearing condition implies that population equals the sum of labor for pro-
duction, entry, research and development, which is
\[ M_t = L_t + \frac{n_t}{\sigma} \left( c_E + \frac{1}{z_R} g_t^t + \frac{1}{2 z_D} \sigma_t^2 \right). \] (11)

Likewise, the capital-market-clearing condition is
\[ K_t = \left( \frac{\alpha}{r_t} \right)^{\alpha+\zeta} \left( \frac{1 - \alpha - \zeta}{w_t} \right)^{1-\alpha-\zeta} N_t \times X_t \] (12)

where \( K_t = k_t M_t \). Finally, the goods-market-clearing condition is
\[ Y_t = C_t + I_t, \] (13)

where \( C_t = c_t M_t \) and \( I_t = K_{t+1} - (1 - \delta)K_t \).

2.7 Equilibrium

We now define the notion of equilibrium in the economy.

**Definition 1.** Given a sequence for labor supply \( \{M_t\} \), an equilibrium is a sequence of prices \( \{w_t, r_t\} \), business choices \( \{l_{i,t}, k_{i,t}, g_t, \sigma_t\} \), household choices \( \{c_t, k_t\} \), a measure of entrants \( \{n_t\} \), and the number of projects, \( \{N_t, S, N_{t, U}, N_t\} \), such that:

- (a) \( c_t \) and \( k_t \) solve the optimization problem of a household (1),
- (b) \( l_{i,t} \) and \( k_{i,t} \) solve the business’s static problem (2),
- (c) \( \sigma_t \) and \( g_t \) are the innovation choices that results from problem (3),
- (d) The free entry condition (4) is satisfied,
- (e) \( N_{t, S}, N_{t, U}, \) and \( N_t \) are in accordance with the laws of motion (8), (9), and (10),
- (f) The clearing conditions for the labor market (11), the capital market (12) and the goods market (13) are satisfied.

2.8 The equilibrium step size of innovation

The optimal step size of innovation is obtained by solving (3). When we incorporate the free entry condition (4) into the solution for the step size of innovation, we find that the step size of innovation is constant in equilibrium. The next lemma presents this result. *All proofs are in appendix A.1.*
Lemma 1 (Step size of innovation). The equilibrium step size of innovation is constant,

\[ g^* = \left( \frac{2cEzR}{\ell - 2} \right)^{\frac{1}{7}}. \tag{14} \]

This lemma implies that the step size of innovation is determined by only three parameters: the slope of the cost of innovation, the entry cost, and research efficiency. The free entry condition is crucial to this result. Many important aspects of the economy influence income levels, but not the size of innovation, as in Atkeson and Burstein (2010). This result simplifies the analysis because it implies that the productivity growth rate of young businesses, determined by \( g^* \), and that of old businesses, determined by \( g_S \), will be constant.

One may think that a model with constant \( g^* \) and \( g_S \) cannot capture what happens in reality. However, it is worth highlighting that Alon et al. (2018), Karahan, Pugsley and Sahin (2019), and Hopenhayn, Neira and Singhania (2022) found that there are minimal changes in the life-cycle profile of business dynamics statistics like the exit rate and average size in the US since there is available data, which is consistent with small changes in \( g^* \) and \( g_S \) over time in our model. For robustness, our quantitative exercises include cases in which \( g^* \) and \( g_S \) are not constant.

3 Balanced growth path

In this section, we characterize a balanced growth path for this economy and investigate how it is impacted by changes in the constant population growth rate \( g_M \). The following lemma characterizes the economy’s BGP equilibrium.

Lemma 2 (Characterization of the balanced growth path). Given a constant growth rate of the labor supply greater than the old businesses’ survival rate, \( g_M > s_{S,\infty} \), there is a unique BGP equilibrium in which the following occurs: (a) Aggregate variables \( Y, K, \) and \( C \) grow at constant rates, (b) Wages grow at the same rate, \( g_w = (gX)(1-\alpha)/\zeta \), (c) The interest rate is fixed at \( r = \frac{(g_w)^{\beta}}{\beta} - (1 - \delta) \), (d) The step size \( g \) and the probability

\(^{11}\)Technically, having a constant step size requires that expected profit be a monomial function of potential productivity \( \hat{x} \).
of starting business \( \sigma \) are constant, (e) Business size is constant because the number of businesses grows at the same rate as the labor in production and population, \( g_N = g_L = g_M \), and (f) Average productivity and successful business productivity grow at the same rate, \( g_X = g_\chi \).

In the BGP described above, the average productivity of all projects, \( X \), is a function of the potential productivity of new projects today, \( \hat{x}_1 \), and other parameters:

\[
X = \frac{\hat{x}_1 \sum_{a=1}^{\infty} \left( \frac{1}{g_X g_N} \right)^{a-1} (\Lambda_{S,a}(g_S) + \Lambda_{U,a})}{\sum_{a=1}^{\infty} \left( \frac{1}{g_N} \right)^{a-1} (\Lambda_{S,a}(1) + \Lambda_{U,a}/\theta)},
\]

(15)

where the growth rate of the number of businesses, \( g_N \), is used to account for the increase in the number of businesses over time. Similarly, \( \chi \), which is today’s reference productivity for innovators, is

\[
\chi = \frac{\hat{x}_1 \sum_{a=1}^{\infty} \left( \frac{1}{g_X} \right)^{a-1} \left( \frac{1}{g_N} \right)^{a-1} \Lambda_{S,a}(g_S)}{\sum_{a=1}^{\infty} \left( \frac{1}{g_N} \right)^{a-1} \Lambda_{S,a}(1)}.
\]

(16)

Because potential productivity today \( \hat{x}_0 = g_X \hat{x}_1 \) equals the step size of innovation multiplied by the average productivity of successful projects; i.e., \( \hat{x}_0 = g_\chi \), we can derive an equation that defines the relationship between the step size \( g \), productivity growth rate \( g_X \), and the number of businesses growth rate \( g_N \). This equation implies that \( g_X \) solves

\[
g = \frac{\sum_{a=1}^{\infty} \left( \frac{1}{g_M} \right)^a \Lambda_{S,a}(1)}{\sum_{a=1}^{\infty} \left( \frac{1}{g_X g_M} \right)^a \Lambda_{S,a}(g_S)}.
\]

(17)

We can immediately see in equation (17) the two sources of growth determining \( g_X \): \( g_S \) and \( g \). We can also see the potential role of population growth \( g_M \), which will be the focus of the next subsection.

To gain more intuition on the workings of the model, consider for a moment a case in which all new businesses become successful at age 1 (\( \lambda = 1 \)). The cost of this simplification is that in this simpler case, we can distinguish only between
entrants and incumbents (not between young and old businesses), and all incumbents’ productivity growth and exit rates will be the same.

In this case, however, we can find a closed-form solution for $g_X$ as a function of $g$, $g_S$, and the share of incumbent businesses. In particular, we find that

$$\text{share of incumbent} = \frac{(s_S/g_M) \times n + (s_S/g_M)^2 \times n + \ldots}{n + (s_S/g_M) \times n + (s_S/g_M)^2 \times n + \ldots} = \frac{s_S}{g_M},$$

and total productivity growth is simply

$$g_X = g_S \times (s_S/g_M) + g^* \times (1 - s_S/g_M). \quad (18)$$

This equation makes it immediately clear that there will be positive total productivity growth ($g_X > 1$) even if only successful businesses have positive productivity growth ($g_S > 1, g = 1$) or if only new businesses have positive innovation ($g > 1, g_S = 1$). In addition, note that this result is regardless of the value of population growth (as long as $g_M > s$).

### 3.1 The “sufficient statistic”

Before studying the impact of population growth on productivity growth, we show that the employment-size growth rate of surviving old businesses converges to the simple ratio of productivity growth rates, $g_S/g_X$.

**Lemma 3** (Growth rate of the size of surviving old businesses). *In a balanced growth equilibrium, the employment growth rate of surviving businesses converges monotonically to $g_S/g_X$ as the age $\rightarrow \infty$.*

As we will show below, this variable will be a “sufficient statistic” for characterizing the influence of population growth on productivity growth.

### 3.2 The impact of population growth on TFP growth

TFP is measured as in the data by

$$\text{TFP} \equiv \frac{Y}{K^{\bar{\alpha}}M^{1-\bar{\alpha}}},$$

17
where the share of capital $\tilde{\alpha}$ is simply

$$\tilde{\alpha} \equiv 1 - \text{labor share of income} = 1 - \frac{wM}{Y}.$$ 

Note that $\tilde{\alpha}$ is different from $\alpha$, and it is constant in a BGP but may vary along a transition and across BGPs.

Starting from the definition of TFP, we can obtain

$$\text{TFP} = \left( \frac{L}{M} \right)^{1-\tilde{\alpha}} \left( \frac{Y}{L} \right)^{\frac{\tilde{\alpha}}{\alpha} - 1} \left( \frac{NX}{L} \right)^{\frac{\tilde{\alpha}}{\alpha} r} \left( \frac{L}{M} \right)^{\frac{(1-\tilde{\alpha})}{1-\alpha}}.$$ 

Given the equality among total labor force growth, labor force for production, and the growth in the number of businesses ($g_M = g_L = g_N$) and constancy of interest rate ($g_r = 0$) along a BGP, growth in TFP along a BGP can be written as

$$g_{\text{TFP}} = g_X \frac{(1-\tilde{\alpha})}{1-\alpha}.$$ 

Therefore, the main question for understanding the impact of population growth on TFP is how population growth affects average productivity growth, or $\frac{dg_X}{dg_M}$. The following lemma, which is the paper’s main theoretical result, employs equation (17) to characterize the impact of $g_M$ on $g_X$.\textsuperscript{12}

**Lemma 4** (The sign of the impact of population growth on productivity growth). In a balanced growth equilibrium, if the growth rate of the size of surviving old businesses is negative, then an increase in the labor force growth rate $g_M$ raises average productivity $g_X$; i.e., if $g_S / g_X < 1 \Rightarrow \frac{dg_X}{dg_M} > 0$.

We explain this result after we present the next result, which describes how the same “sufficient statistic” determines the size of the impact of population growth on productivity growth.

\textsuperscript{12}As the share of capital $\tilde{\alpha}$ varies across BGPs, $d\tilde{\alpha}/dg_M$ is a non-zero value. Specifically, it is most likely negative, thereby amplifying the effect of population growth on TFP growth. Further discussion on this topic can be found in appendix A.2.
Lemma 5 (Magnitude of the impact of population growth on productivity growth).

Suppose the growth rate of the size of surviving old businesses is negative. Suppose there are two economies with the same average productivity growth (\(g_X\)) and the same labor force growth (\(g_M\)), but the growth rate in the size of old businesses decreases faster in one than in the other.\(^{13}\) Then the impact of population growth on productivity growth, \(\frac{dg_X}{dg_M}\), is larger in the economy in which the growth rate in the size of old businesses decreases faster.

These findings are better understood by considering that there are two mechanisms. First, recall that the growth in the number of businesses must be equal to the growth rate in the number of new businesses. Therefore, an increase in \(g_M\) (and consequently in \(g_N\)) reduces the share of old businesses. Second, total productivity growth equals the weighted average of productivity growth of businesses of various ages. As a result, if the productivity growth of old businesses is lower than the average, a decline in \(g_M\) (and an increase in the share of old businesses) will have a negative effect on total productivity growth. Because the ratio of the average productivity growth rate to the productivity growth rate of old businesses equals the growth rate of the size of surviving old businesses, we referred to this variable as a "sufficient statistic," as it is the only information required to identify the sign of the impact of population growth on productivity growth.

To see this logic more clearly in an equation, recall the case in which all new businesses become successful at age 1 (\(\lambda = 1\)). There, total productivity growth is given by equation (18). Clearly, productivity growth (\(g_X\)) is the weighted average of incumbent businesses’ productivity growth (\(g_S\)) and the step size of innovation by new businesses (\(g\)). Decreasing population growth increases the weight on incumbents (\(s/g_M\)), and it has a negative impact on total productivity growth (\(g_X\)) as long as the size of incumbent businesses declines over their life-cycle (what hap-

\(^{13}\)Although productivity growth is endogenously determined in this model, we can set productivity growth arbitrarily by adjusting \(z_R\).
pens if \( g_S / g_X \)). These findings also imply that the calibration of the productivity life-cycle profile is crucial for the quantitative results presented in the next sections of this paper.

This section’s results are achievable because there is a perfectly elastic supply of new businesses, entrants’ innovation is constant, and we assume that survival rates and the growth rate of productivity of successful businesses are exogenous. In the following sections, we study versions of the model in which we allow for congestion at entry, non-constant entrants’ innovation, and endogenous exit and innovation by successful businesses. Our quantitative analysis reveals that these forces amplify the effects discussed in this section.

4 Quantitative model

4.1 Entry congestion and innovation spillovers

Before proceeding to the quantitative analysis of the model, we add two realistic features to the model presented in the previous section: congestion of entering businesses and spillovers to older businesses. These extensions will depend on two key parameters: \( \phi \) and \( \gamma \). By setting \( \phi = \gamma = 0 \), these two extensions can be removed. This exercise will assess the importance of these features in terms of our quantitative results.

First, we modify the free entry condition (4) to account for potential “congestion.” According to Hopenhayn (1992), the working assumption in the model described above is that as long as the free entry condition is satisfied, the number of entrants is perfectly elastic, so \( n_t \) can be chosen to scale up or down the number of businesses and clear the labor market. We modify the free entry as in Karahan, Pugsley and Sahin (2019) to add a more realistic response of the number of businesses to economic conditions. In particular, we replace \( c_E \) with \( c_E(\bar{n}_M)\phi \), where \( \bar{n} = n_t / n_i \) and \( \bar{n} \) represents the number of entrants in normal times (i.e., our reference period 1980-1999), and \( \bar{M} \) is defined analogously. Now, the modified
free entry condition is \( V_t \leq w_t c_E(\tilde{n}_t/\tilde{M}_t)^\phi \), which means that in order to increase the number of entrants into the economy, the value of entry must also increase. The key parameter is \( \phi \), which we will calibrate based on previous estimates.

The main implication of this feature is that the step size of innovation is no longer independent of the growth rate of the population. Because of congestion, equation (14) is replaced by \( g^* = \left( \frac{2c_E(\tilde{n}/\tilde{M})^\phi z_R}{t-2} \right)^{1/2} \). Now, the share of entry in the population, \( n/M \), affects the innovation intensity, \( g^* \). Therefore, there is another channel through which the population growth rate, \( g_M \), affects the economy’s productivity growth \( g_X \).

Second, we consider that the productivity growth of successful projects, \( g_S \), may be a function of the last previous productivity growth of successful businesses. This equation captures the idea that successful businesses may benefit (with some delay) from younger businesses’ innovation. Thus, the productivity growth rate of already-successful businesses is

\[
g_{S_t} = \bar{g}_S + \gamma(g_{X_t-1} - \bar{g}_X),
\]

where \( \bar{g}_S \) is a constant representing the productivity growth rate of successful businesses in normal times and \( \bar{g}_X \) is a constant growth rate for successful businesses’ productivity in normal times. The key parameter is \( \gamma \), which we will estimate using data on the relationship between employment growth by mature businesses and overall productivity growth (more on this in the calibration section).

Since the productivity growth of successful projects \( g_S \) depends on \( g_X \), which is equal to \( g_X \) along a balanced growth path, expected productivity for successful projects \( \Lambda_{S,a}(g_S) \) will depend on \( g_X \); i.e., \( \Lambda_{S,a}(g_S(g_X)) \). Thus, equation (17), which determines the growth rate of productivity in the economy, is replaced by

\[
\left( \frac{2c_E z_R}{t-2} \right)^{1/7} = \frac{\sum_{a=2}^{\infty} \left( \frac{1}{\bar{g}_X} \right)^{a-1} \Lambda_{S,a}(1)}{\sum_{a=2}^{\infty} \left( \frac{1}{\bar{g}_X \bar{g}^N} \right)^{a-1} \Lambda_{S,a}(g_S(g_X))}.
\]

If spillover is positive, this extension will amplify the effect of \( g_M \) on \( g_X \).
Although this will be our benchmark quantitative model for the rest of the paper, at the end of this section, we add two more features to the BGP analysis: endogenous exit and endogenous choice of innovation by successful businesses. Incorporating these features substantially complicates the model, but their effect is limited, and if anything, they amplify the results implied by our benchmark model.

4.2 Calibration

We calibrate the model to aggregate statistics and business dynamics data for the United States and Japan. The model can reproduce key stylized facts with relatively few parameters. What is important for our results is not the micro-level information about firm dynamics but aggregate moments about business dynamics by age.\textsuperscript{15} In particular, the model should reproduce the average size of all age-\textit{a} businesses (defined as total employment over total number of age-\textit{a} businesses) and the transition rates from age \textit{a} to age \textit{a} + 1, which will be given by the percent of age-\textit{a} businesses that exit (i.e., the exit rate) and the change in the average size of businesses between age \textit{a} and \textit{a} + 1 restricted to businesses that are in operation at both ages \textit{a} and \textit{a} + 1 (i.e., the growth rate of the size of surviving businesses).

We calibrate the model to reproduce the average from 1980 to 1999. The calibrated parameters based on previous papers or obtained directly from data are shown in the top panel of Table 1. We assign values to the remaining parameters, which are shown in the bottom panel of Table 1, in order to reproduce key stylized facts such as establishment size, life-cycle profiles, and exit rates.\textsuperscript{16}

Thus, the value of \( g_M \) in the US, 1.0143, and in Japan, 1.0103, are the averages for the years 1980-1999. The exponent of the research cost function, \( \iota \), is set

\textsuperscript{15}We could add more heterogeneity by adding iid productivity shocks, and it would not change the results as long as we reproduce business dynamics by age.

\textsuperscript{16}We use establishment-level data to capture product or project-level activity, which is more in line with the process of innovation that we model. Establishment-level data are frequently used as a proxy of project-level analysis, assuming that one establishment produces one product (e.g., Klenow and Li (2020) and Garcia-Macia, Hsieh and Klenow (2019)).
Table 1: Parameters’ values and targets of calibration

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Value</th>
<th>Basis</th>
</tr>
</thead>
<tbody>
<tr>
<td>Entry cost, $c_E$</td>
<td>1</td>
<td>Normalization</td>
</tr>
<tr>
<td>Decreasing returns, $\zeta$</td>
<td>0.2</td>
<td>Standard</td>
</tr>
<tr>
<td>Capital share, $\alpha$</td>
<td>0.32</td>
<td>Standard</td>
</tr>
<tr>
<td>Depreciation rate, $\delta$</td>
<td>0.07</td>
<td>Standard</td>
</tr>
<tr>
<td>Risk aversion, $\epsilon$</td>
<td>2</td>
<td>Standard</td>
</tr>
<tr>
<td>Discount factor, $\beta$</td>
<td>0.96</td>
<td>Standard</td>
</tr>
<tr>
<td>Labor force growth rate, $g_M$</td>
<td>(1.0143, 1.0103)</td>
<td>Average $g_M$ 1980-1999</td>
</tr>
<tr>
<td>Research cost exponent, $\iota$</td>
<td>2.56</td>
<td>GHS</td>
</tr>
<tr>
<td>Convexity of aggregate entry cost, $\phi$</td>
<td>0.55</td>
<td>KPS</td>
</tr>
<tr>
<td>Elasticity of $g_S$ to $g_M$, $\gamma$</td>
<td>0.342</td>
<td>See appendix A.3.</td>
</tr>
<tr>
<td>Research cost slope, $z_R$</td>
<td>(0.933, 1.762)</td>
<td>Average prod. growth</td>
</tr>
<tr>
<td>Development cost slope, $z_D$</td>
<td>(2.413, 1.417)</td>
<td>Average estab. size</td>
</tr>
<tr>
<td>Jump of prod. at success, $1/\theta$</td>
<td>(16.5, 28.1)</td>
<td>Average size by age</td>
</tr>
<tr>
<td>Success probability, $\lambda_a$</td>
<td>See Figure 9</td>
<td>Growth of estab.</td>
</tr>
<tr>
<td>Productivity growth of successful estab., $g_S$</td>
<td>(1.054, 1.023)</td>
<td>Growth of old estab.</td>
</tr>
<tr>
<td>Survival of successful estab., $s_S$</td>
<td>(0.965, 0.973)</td>
<td>Exit rate of old estab.</td>
</tr>
<tr>
<td>Survival of unsuccessful estab., $s_{U,a}$</td>
<td>See Figure 9</td>
<td>Life-cycle profile of exit rate</td>
</tr>
</tbody>
</table>

Note: The parameters with different values for the United States and Japan are shown in parenthesis, with the United States representing the first number and Japan representing the second. GHS is an abbreviation for Greenwood, Han and Sanchez (2022), and KPS is an abbreviation for Karahan, Pugsley and Sahin (2019).

to the same value as in Greenwood, Han and Sanchez (2022). They estimated it to be equivalent to the impact of innovation expenditures on a firm stock market value. Similarly, the parameter $\phi$ that determines the degree of congestion is set as in Karahan, Pugsley and Sahin (2019). The parameter $\gamma$, which influences the diffusion of innovation from new to old businesses, is calibrated using the value estimated in appendix A.3. There, we compare various specifications for regressing current old-establishment productivity growth on the economy’s past productivity growth.
We now discuss the rationale behind the selection of the parameters that match specific targets. To start, it should be mentioned that since all parameters’ values are determined during the process of matching moments, there isn’t a one-to-one link between parameters and targets.

First, because the value of productivity for research, $z_R$, affects the growth rate of average productivity in the economy, we calibrate the BGP to replicate the average productivity growth in the United States and Japan from 1980 to 1999.

Second, note that the jump in productivity when businesses become successful, governed by $\theta$, is a normalization.\textsuperscript{17} We choose $\theta$ so that successful businesses employ, on average, around 80 workers, whereas unsuccessful businesses employ, on average, about 5 workers. Then, we calibrate the success probability $\lambda_a$ for ages $a \geq 2$ by assuming it changes exponentially with age in order to minimize the number of parameters to search on. Thus, all that is required are the starting probability and the decay constant. Figure 9 shows the resulting life-cycle profile of success probability. Both the United States and Japan have extremely low success probabilities. As a result, successful businesses are uncommon. In addition, note that we find that $\lambda_a$ is decreasing in age $a$, which is consistent with Greenwood, Han and Sanchez (2022)’s finding that the odds of success by venture capital funding round decreases with the age of the project.

Third, old businesses’ employment growth is $g_S/g_X$ as discussed in Lemma 3. As a result, we calibrate $g_S$ so that the model accurately reproduces the employment growth for old establishments in the data.

Finally, we calibrate the parameters that determine the survival probability lifecycle profile. We assume that the probability of successful businesses surviving is constant. The resulting survival probability, 0.97, is very similar for the United States and Japan and reflects the fact that the exit rate of old establishments is quite

\textsuperscript{17} The success probability would decrease if we increase the value of $\theta$. In other words, if successful businesses are larger, the calibration would indicate that there are fewer of them to match the growth rate of surviving businesses. This trade-off has a limited impact on the results: as shown in Table 7, the effect of the change in $\lambda$ and $\theta$ are relatively small in opposite directions.
low in both countries’ data. Because successful businesses are rare in the model, the resulting life-cycle profile of survival probabilities for unsuccessful businesses is closely related to the survival probability for establishments in the data.\textsuperscript{18} The resulting profiles for $s_S$ and $s_U$ for the United States and Japan are also shown in Figure 9 in appendix A.4.\textsuperscript{19}

As a result, we calibrate 10 parameters using 32 moments (31 bars in Figure 3 and average productivity growth) for the US and 20 moments (19 bars and average productivity growth) for Japan. They are targeted with an equal weight.

Figure 3 depicts the fit of calibration targets.\textsuperscript{20} Figure 3’s left panel shows how well the model matches the exit rate life-cycle patterns for the US and Japan. Exit rates in both countries fall with age, and exit rates in the US are greater than in Japan, especially for newer establishments. The model accurately predicts these trends, which is critical for determining the relative relevance of young and old businesses.

The middle panel of Figure 3 depicts the fit of the life-cycle profile of establishment size as measured in employment. For the plot, we normalize employment by the employment of age-one establishments such that the plot starts at one. Two facts are important. First, the model accurately reproduces the profiles for the US and Japan. Second, Japan’s profile is much lower than that of the US. While establishments 27 years or older in the US are approximately 3.5 times larger than those one-year-old, the same ratio for establishments 29 years or older in Japan is 1.5.

\footnote{In the same manner as the choice of success probability, we calibrate the survival probability $s_{U,a}$ using an exponential decay function while allowing it to converge to a non-zero value. Hence, the initial probability, long-run probability, and decay constant are calibrated.}

\footnote{We show in Figure 8 that the model also reproduces well the share of young businesses, which could be expected given the fit of the exit rates.}

\footnote{The source of data for the United States is the Business Dynamics Statistics (BDS), which the United States Census Bureau produces. In Figure 3, the average of 1980-2019 is taken for all three measures. The data source for Japan is the Economics Census and Establishment and Enterprise Census conducted by the Statistics Bureau. The exit rate and growth of surviving establishments by age are based on data from 2004, as these measures are only available in 2004 compared with the data from 2001. For the employment size, we have extracted the year-of-birth fixed effect as the life-cycle profile for Japan, unlike the one for the US, is influenced by the year of birth. More on
Finally, Figure 3’s right panel shows the growth of surviving businesses. Contrary to the average establishment size by age, these profiles are unaffected by business selection into exit. Thus, the differences between the middle panel and the right panel help identify the differences in survival rates of successful and unsuccessful businesses.\footnote{In particular, if the growth rate of average employment size is equal to the growth rate of average employment size of surviving businesses, then the exit rate of successful and unsuccessful businesses should be equal.} As our theoretical analysis showed, it is very important to reproduce the growth of surviving businesses. Critically, there is a diminishing growth profile with respect to age, and old businesses experience a decline in size on average. Recall this was our “sufficient statistic” described in Lemma 3. The patterns in the data indicate that a drop in population growth will lead to a fall in aggregate productivity growth along the BGP, as stated by our Lemma 4. Also, the decline in the size of surviving old businesses is faster in Japan than in the US. Given the finding in our Lemma 5, we are likely to find a larger impact of labor force growth on productivity growth in Japan than in the US. In terms of the cali-

\textbf{Figure 3: Fit of life-cycle profiles}

...
bration, the growth rate of the size of surviving old businesses is the prime target for calibrating $g_S$. Consequently, we obtain $g_S = 1.054$ for the US and $g_S = 1.023$ for Japan.

5 Balanced growth path analysis

This section shows how variations in the population growth rate impact the productivity growth rate along the BGP.

5.1 Results using the benchmark model

We first offer the benchmark findings before showing how spillovers and congestion affect the results. The exercise is straightforward. We take the model’s BGP calibrations for the US and Japan from 1980 to 1999, modify $g_M$, and find the new value of $g_X$, which is proportional to $g_{TFP}$ in the BGP. This exercise is repeated for various values of $g_M$. Figure 4 shows the results for the US (left panel) and Japan (right panel).

Figure 4: Impact of population growth on TFP (comparison across BGPs)

The calibrated point is indicated by the stars in Figure 4, and the lines depict how $g_{TFP}$ changes as $g_M$ increases. Additionally, vertical lines represent historical times of high labor force growth and the labor force growth projections for the
years 2050 to 2060. Using these time frames as examples, we illustrate how drops in labor force growth imply major shifts in the pace of productivity growth of the economy. Our model predicts a 0.3-percentage-point drop in productivity growth for the US as a result. For Japan, population growth declined by more than 3 percentage points between the indicated 100 years (1950-1960 to 2050-2060), which implies a 0.6 percentage point reduction in productivity growth.

5.2 Role of key model’s features

What role do the key characteristics of the model play in the outcomes shown in Figure 4? To address this, we compare our benchmark model (column A) results with those for three alternative models in Table 2. The findings of a model with no congestion ($\phi = 0$), a model with no spillovers ($\gamma = 0$), and the simplest model with neither congestion nor spillovers ($\phi = \gamma = 0$) are shown in columns B, C, and D, respectively.

The results in Table 2 show that congestion and spillovers are important, although around 65% of the impact would still be seen absent them. The overall impact for the US decreases from 0.30 percentage points to 0.19 percentage points when both factors are removed. Japan’s decrease falls from 0.60 to 0.39 percentage points. As a consequence, we conclude that the impact is a reduction of about 0.08-0.12 percentage points of productivity growth for every percentage point reduction in the US population growth and between 0.14 and 0.21 percentage points of productivity growth for every percentage point reduction in Japan’s population growth.

Why do these features amplify the impact on TFP growth? First, congestion reduces the entry cost when the population growth rate decreases because there are fewer entrants. This declining entry cost discourages innovation efforts, as entrants can compensate for the entry cost with lower productivity. Consequently,

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22 The labor force growth projections are taken from the Bureau of Labor Statistics (BLS) for the US and the Cabinet Office (CAO) for Japan.
Table 2: The role of key features on the impact of population growth on TFP along the BGP

<table>
<thead>
<tr>
<th>Periods</th>
<th>United States</th>
<th>Japan</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Data’s growth in labor force, %</td>
<td>Model’s implied growth in TFP in the BGP, %</td>
</tr>
<tr>
<td>United States</td>
<td></td>
<td></td>
</tr>
<tr>
<td>1900-1910</td>
<td>2.60</td>
<td>1.35</td>
</tr>
<tr>
<td>1980-1999</td>
<td>1.43</td>
<td>1.20</td>
</tr>
<tr>
<td>2050-2060</td>
<td>0.25</td>
<td>1.05</td>
</tr>
<tr>
<td>Difference in pp</td>
<td>-2.35</td>
<td>-0.30</td>
</tr>
</tbody>
</table>

| Japan         |               |              | (A)         | (B)        | (C)         | (D)       |
| 1950-1960     | 1.94          | 1.08         | 1.05        | 1.03       | 1.01        |
| 1980-1999     | 1.03          | 0.89         | 0.89        | 0.89       | 0.89        |
| 2050-2060     | -0.95         | 0.48         | 0.53        | 0.58       | 0.62        |
| Difference in pp | -2.89        | -0.60        | -0.52       | -0.45      | -0.39       |

Congestion increases the effect through smaller innovation. Second, the impact of spillovers is more straightforward: the decline in population growth leads to lower productivity growth, resulting in smaller spillovers to the productivity growth of existing businesses, which is also part of overall productivity growth.

In addition, Table 7 in online appendix B.3 presents the results of a sensitivity analysis for the size of the impact of $g_M$ on $g_{TFP}$. We find that the most important parameters affecting the size of the impact are the survival of successful businesses, $s_S$, and their productivity growth, $g_S$. We find that an increase in $s_S$ or a decrease in $g_S$ would increase the size of the impact of $g_M$ on $g_{TFP}$ significantly. Note that the last result is in line with Lemma 5.

5.3 Incorporating endogenous exit

In this section, we consider an extension of the model incorporating endogenous exit. To give businesses a reason to exit, we incorporate a fixed cost shock $c_{awe}$ to the profits of unsuccessful businesses that depend on the age of the busi-
ness.\textsuperscript{23} If the pre-fixed costs expected discounted profits $I_a$ is larger than the fixed cost, they pay the fixed costs and continue their businesses. If not, they exit the market.

We also keep an exogenous exit probability at the exit rate level for very old unsuccessful businesses in the original calibration, $s_{U,\infty}$. Thus, the survival probability for unsuccessful businesses is now given by $s_{U,a} = s_{U,\infty} \Pr(c_a \omega \epsilon \leq I_a) = s_{U,\infty} \times F(I_a / (w_c a))$, where $F$ is the distribution of $\epsilon$. For tractability, we assume that $F$ is Type III extreme value (or Weibull) distribution; i.e., $\epsilon \sim \text{Weibull}(1, \vartheta)$. With this assumption, the survival probability for unsuccessful businesses is given by

$$s_{U,a} = s_{U,\infty} \left[ 1 - \exp \left( - \left( \frac{I_a}{w_c a} \right)^{\vartheta} \right) \right],$$

where $\vartheta$ is a parameter of the distribution. In addition, we incorporate the expected fixed cost in expected profits given survival. We calibrated $c_a$ to get the same age profile of the exit rate as the exogenous exit case. Also, we chose $\vartheta = 0.74$ such that the effect of population growth on the economy’s exit rate across BGPs coincides with the effect in Hopenhayn, Neira and Singhania (2022).\textsuperscript{24}

Figure 5 shows how the effect of the population growth rate on TFP growth compares between the model introduced in section 2 and the model with endogenous exit. The effect is very small for the US and Japan’s calibrations. Why is the amplification so small? Because there are two counterbalancing forces. First, a decline in population growth produces an increase in survival rate for each age because wages grow slower. This effect on the survival rate for every age increases the share of old businesses, which magnifies the effect on productivity growth due

\textsuperscript{23}Successful businesses could also have to pay a fixed cost, but we assume it is negligible compared to their revenue, so they would not exit for this reason (they are still subject to the exogenous probability of survival).

\textsuperscript{24}In particular, we match that when population growth declines from 2.66\% to 0.78\%, the exit rate declines by 0.88 percentage points. As a reference, with exogenous exit, the composition effect due to firms getting older would imply a decline of 0.72 percentage points.
to the growth composition effect. Second, there is also a change in the step size of innovation, $g$. Given that the survival rates increase endogenously for each age, businesses find it beneficial to innovate more.

**Figure 5: Endogenous exit and successful businesses innovation**

5.4 Incorporating endogenous innovation by successful businesses

In this section, we incorporate endogenous productivity growth by successful businesses, $g_S$ into the model.

Denote $I_a(x)$ be the value of a successful business at age $a$. It satisfies the following Bellman equation:

$$I_a(x, X, w) = S(x, X, w) + \frac{s}{1 + r - \delta} \max_{g_S} \left[ I_{a+1}(g_S x, g_X X, g_w w) - C_S(g_S) w \right], \tag{20}$$

where $C_S(g_S)$ is the cost function of achieving productivity growth of $g_S$. We consider a general case for $C_S(g_S)$ that allows for the cost of productivity growth to depend on the distance of the business productivity to the frontier (i.e., the average of successful businesses, $\chi$) and to the average productivity of the economy (i.e., $X$). In particular, we assume

$$C_S(g_S) = c_S g_S \left( \frac{x}{\chi} \right)^{\xi} \left( \frac{x}{X} \right)^{1-\xi}, \tag{21}$$
and let $\xi$ determine which distance is more relevant. In this case, we find that the endogenous innovation by successful businesses is

$$\begin{align*}
g_S &= \left( \frac{\chi}{\lambda} \right)^{\xi} \left( \frac{c_S}{D} \right)^{\frac{1}{t^2}} - \theta \frac{1}{\lambda \chi} \left( \frac{\chi}{\lambda} \right)^{1-\xi} \frac{\lambda}{1-\theta} \left( \sum_{a=1}^{\infty} \beta a \left( \frac{g_w}{g_X} \right)^{a} \Lambda_{u,a} \right) \right)^{\frac{1}{1-\xi}},
\end{align*}$$

where

$$D = z^{\frac{1}{2}} \left( \frac{z_R}{l} \right)^{\frac{2}{t^2}} \left( \frac{t - 2}{2l} \right)$$

is a constant value. To obtain quantitative results, we calibrate two parameters: $c_S$ and $\xi$. We calibrate $c_S$ to have the same $g_S$ in the reference period (1980-1999) as in our benchmark quantitative model, and $\xi$ to have the elasticity of $g_S$ to $g_X$ consistent with the result in Table 6 when comparing between the averages for 1980-1999 and 2000-2019. We find that innovation by successful businesses is hump-shaped in population growth; as a consequence, the effect of population growth on TFP growth is larger for the low-population growth states and smaller for the high-population growth states compared with the simplest case. This implies that the effect on TFP growth can be larger in the future than the main analysis, as the population growth is slowing down.

6 Transitional Dynamics

The concern that declining population growth will have an influence on TFP growth is clearly long-term. However, we address several interesting questions by computing transitional dynamics. In this section, we compute transitional dynamics and use the results to answer three questions. The first question is: How important has population growth been for the slowdown in TFP in the US and Japan in the last 40 years? We show that the share accounted for population growth is significant, but it is a fraction of the differences among BGPs. This result suggests that the transition is slow, and a significant part of the impact of the decline in population growth on TFP growth will occur in the future. Thus, a natural question is:
How much of this effect will be observed in the future? Finally, we ask: What are the causes of the TFP growth’s sluggish reaction to population growth changes? We show that the interplay of level-vs-growth and labor-reallocation effects is partially responsible for this.

6.1 Impact on TFP growth in the last 40 years

We use transitions between BGPs to compare the drop caused by the decline in population growth with data for the United States and Japan.

We select as the starting BGPs those that correspond to labor force growth in the first years that we observe a significant decline in the trend of the growth rate of the labor force (1950 for Japan and 1970 for the US).25 We utilize the BGPs corresponding to the value of $g_M$ in 2020 for the final BGP in both countries.26 The entire transitions for the main variables are shown in Figure 10 in the online appendix.

Table 3 describes the changes in the last 40 years, focusing on two sub-periods.27 respectively. The numbers for the “slowdown” in TFP growth are the difference in TFP growth (in percentage points) between the average trend growth in 1980-99 and 2000-19. For instance, TFP trend growth in the US was 1.195% in 1980-99 and 1.010% in 2000-19, so the difference is 0.185 pp. The decline in TFP growth was more severe in Japan: on average, it was 0.763% in 1980-99 and 0.345% in 2000-19, so the difference is 0.418 pp.

How much of the drop in TFP growth can be attributed to the slowdown in population growth? To answer this question, we take the simulated transitions

---

25 The Christiano and Fitzgerald (2003) filter is used to extract the slow-moving trend of labor force growth. We chose this filter because it allows us to choose the parameters to capture the long-run movement in the labor force. We set the parameters at 2 and 40 and also consider the changes using other values.

26 In a later section, we show the amplification obtained by including forecasts for $g_M$ until 2060.

27 We begin the analysis in 1980 since other circumstances, such as World War II and the high economic growth period in Japan in the post-war period would likely be influential in previous decades. Recall that in the previous section, we calibrated the model to have a balanced growth rate of TFP equal to the average growth of TFP between 1980 and 1999. To facilitate the comparison with data, for the transitions, we recalibrated $z_R$ for the US and Japan to make sure that the average growth in the period 1980-1999 is equal in the model in the data. Only very small changes were necessary; the new values are 0.886 and 1.622 for the US and Japan.
Table 3: TFP growth slowdown since 1980-99 to 2000-19, data-model comparison

<table>
<thead>
<tr>
<th></th>
<th>United States</th>
<th>Japan</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Change in $g_{TFP}$</td>
<td>Share accounted for</td>
</tr>
<tr>
<td>Data</td>
<td>0.185</td>
<td>–</td>
</tr>
<tr>
<td>Benchmark</td>
<td>0.088</td>
<td>47.5%</td>
</tr>
<tr>
<td>No congestion</td>
<td>0.078</td>
<td>42.0%</td>
</tr>
<tr>
<td>No spillover</td>
<td>0.073</td>
<td>39.3%</td>
</tr>
<tr>
<td>Simplest</td>
<td>0.067</td>
<td>36.2%</td>
</tr>
</tbody>
</table>

given the evolution of the trend in population growth and compute the average growth in TFP in the period 1980-1999 and 2000-2019 as we did in the data. Using the benchmark model, we find that the drop in population growth accounts for 47.5% of the decline in TFP growth in the US. If we abstract away from congestion, this share is reduced by about one-tenth, while removing spillovers decreases it by less than one-fifth. In the simplest scenario, the drop in population growth explains 36.2% of the decline in TFP growth in the US over that time period. The analysis is similar in Japan. The benchmark model accounts for roughly 17.4% of the drop in TFP growth, whereas the simplest model (without congestion and spillovers) accounts for around 11%.

6.2 Future decline in productivity

The results in Table 3 together with BGP analysis shown in Figure 4 suggest that the transition between BGPs is slow and, as a consequence, part of the effect of population growth on productivity growth will occur in the future. The top panel of Table 4 shows the expected decline in TFP growth between 2020 and 2050 and between 2020 and 2100. The expected changes for these subperiods in the US are -0.05 and -0.05 percentage points, respectively. Note that this implies that the impact of population growth on productivity growth will increase by almost 50% in the future (it was -0.088 between the average of 1980-1999 and 2000-2019). The
expected change in productivity in Japan is much larger: -0.19 and -0.24 percentage points between 2020 and 2050 and between 2020 and 2100, respectively. The faster labor force growth decline in Japan than in the US between 2005 and 2020 causes this result.

To complete the analysis of expected TFP growth, we re-computed the transitions, incorporating the forecast for the decline in labor force growth until 2060. The bottom panel of Table 4 shows the results. Since labor force growth is expected to continue decreasing, the impact is larger, particularly between 2020 and 2100. For that period, the expected change in productivity growth is -0.08 for the US and -0.35 for Japan.

Table 4: Expected change in TFP growth as a consequence of the decline in the growth of the labor force (percentage points)

<table>
<thead>
<tr>
<th>Country</th>
<th>United States</th>
<th>Japan</th>
</tr>
</thead>
<tbody>
<tr>
<td>Benchmark</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Between 2020 and 2050</td>
<td>-0.05</td>
<td>-0.19</td>
</tr>
<tr>
<td>Between 2020 and 2100</td>
<td>-0.05</td>
<td>-0.24</td>
</tr>
<tr>
<td>Including forecast for $g_M$</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Between 2020 and 2050</td>
<td>-0.05</td>
<td>-0.27</td>
</tr>
<tr>
<td>Between 2020 and 2100</td>
<td>-0.08</td>
<td>-0.35</td>
</tr>
</tbody>
</table>

6.3 Why is the response so slow? Two counterbalancing factors

Why is there a slow reaction of $g_{TFP}$ to $g_M$, as mentioned in the preceding subsection? To answer this question, we evaluate the economy’s response to a one-time change in population growth. We carry out this experiment in the model calibrated for the US.
Recall that TFP is given by

\[
TFP = \left( \frac{L}{M} \right)^{1-\tilde{\alpha}} \left( \frac{\alpha}{r} \right)^{\frac{\tilde{\alpha}}{1-\tilde{\alpha}}} \left( \frac{NX}{L} \right)^{\frac{1}{1-\alpha}}.
\]

So, in a transition, TFP growth is not just equal to growth in average productivity $X$. As long as the share of workers in the production sector, $L/M$, the interest rate, $r$, and the average firm size, $N/L$, are changing, they will cause changes in TFP growth. These three drivers of TFP growth in the transition are absent in the BGP because they are constant. It turned out that the most important of these three components is $L/M$, which should not be surprising because it has the largest exponent.

Figure 6 depicts the economy’s response to the change in population growth rates that we feed into the model. The shock is a permanent fall in population growth from 2% to 1% in the year designated as zero.\textsuperscript{28} The values displayed are relative to the values in the BGP with population growth equal to 2%.

First, note that in the period the shock is realized, population growth (top left panel) is half of the original value, and the growth in the number of businesses (top middle panel) also falls quickly, although a bit more slowly. The decline in the growth rate of the number of businesses generates a slow fall in the share of young businesses (top right panel), which takes more than 20 years to reach the value of the new BGP. As discussed previously, in the long run, this generates a decline in average productivity growth (bottom left panel). However, during the shock period and for a few years afterward, average productivity growth is larger than in the original BGP. This is referred to as the level-vs-growth effect. The long-run decline in average productivity growth is due to a rise in the proportion of older businesses in the economy, which have lower productivity growth than the average. However, on impact, a higher proportion of older businesses

\textsuperscript{28}For simplicity, we assume agents learn about the shock only one period before it occurs. This assumption allows for the growth in a number of businesses to fall in the same period as population growth, as the decision to enter is made one period in advance.
positively influence average productivity growth because younger businesses are less productive than older businesses. Thus, in the short run, younger businesses’ lower productivity level outweighs their greater productivity growth, and average growth productivity rises. The decline in the entry rate captured by the drop in the growth rate of the number of businesses (shown in the top middle panel) generates a decline in employment in the innovation sector of the economy, so a larger share of workers are employed in the production sector (L/M increases in the bottom middle panel). This effect, which we refer to as the labor-reallocation effect, is the second driver of the short-run increase in TFP growth after a decline in the growth rate of the population.

The labor-reallocation and level-vs-growth effects operate together to drive TFP growth (panel F) in the short run in the opposite direction that it moves in the long run as population growth changes. As a result, these forces partially counterbalance the short-run effect on TFP growth and create its sluggish response to changes in population growth.
We also study the sensitivity of the speed of convergence to the model’s parameters in Table 7 in the online appendix B.3. The most important parameters are $s_S$, $s_U$, $g_S$, and $\beta$. We find that declines in $s_S$, $s_U$, $g_S$, and $\beta$ would significantly increase convergence speed.

7 Validation using US state-level data

As mentioned in the introduction, there is vast empirical literature documenting the impact of population growth on productivity growth. In this section, we offer new evidence using state-level US data. The purpose is to validate the proposed mechanism as closely as possible.

First, we study the impact of population growth on productivity growth throughout US states using local projections. The focus is on studying if the dynamics after a change in population growth described in the previous section are present in the data. Second, we consider the possibility of endogeneity in the regressions using an instrumental variable approach. The results in the instrumental variable approach resemble those in the natural experiment studied by Peters (2022).

For the analysis in this section, we would ideally need a lengthy time series of state-level TFPs, which are not available in the US. As a result, we use real GDP per worker, which is referred to as labor productivity. Our real GDP per worker and labor force data range from 1977 to 2019. To keep the analysis comparable, we also use labor productivity from the model, which is calculated using the following expression:

$$\log(g_{prod,t}) = \frac{1}{1 - \tilde{\alpha}} \times \log(g_{TFP,t}) - \frac{\tilde{\alpha}}{1 - \tilde{\alpha}} \times \log(g_{r,t}).$$

We run an anticipated transition from 1900 to 2060 for the ten largest US states to generate a simulated time series from the model.\footnote{See Peters (2022) and references therein.}

\footnote{A concern about using state-level data is that there may be knowledge spillover across states. Note that if the spillover is positive, meaning that faster growth in one state generates higher growth in other states, then the coefficient in our regression will be downward biased.}

\footnote{Not surprisingly, the estimates are very comparable if we use simulated aggregate US data or the simulation for only two states.}
7.1 Local projections

A difficulty with the analysis of the relationship between changes in productivity growth and labor force growth is that, as shown in the previous section, it is not monotonic. Productivity growth rises at first and then falls due to a slowdown in labor force growth, as shown in Figure 6. This pattern suggests that we should estimate a dynamic model. As a result, we follow Jordà (2005) and employ local projections.

Our left-hand-side variable is the change in labor productivity growth rate between \( i \) years after the shock and the year before the shock, for each state \( s \),

\[
\Delta (g_{prod})_{s}^{i,t} = g_{prod,t+i}^{s} - g_{prod,t-1}^{s}.
\]

We regress this variable on the change in the growth rate of the labor force,

\[
\Delta (g_{prod})_{s}^{i,t-1} = \beta_{0}^{i} + \beta_{1}^{i} \times \Delta (g_{M})_{s,t-1}^{i} + \text{controls},
\]

for \( i = 0, 1, ..., 6 \). Note that these are seven different regressions, one for each value of \( i \). The controls included are four lags of \( \Delta (g_{M})_{s,t-1}^{i} \), four lags of \( \Delta (g_{prod})_{s,t-1}^{i} \), and a quadratic polynomial in year.

Figure 7: Change in labor productivity growth after a 1pp decline in labor force growth

![Figure 7: Change in labor productivity growth after a 1pp decline in labor force growth](image)

Note: The shaded areas represent one (darker) and two (lighter) standard error bands.

Figure 7 depicts the outcome of estimating local projections in data for US states. The shape of the response resembles the results generated with the model.
in Figure 6. A reduction in labor force growth has a positive (0.137) and significant (at 1%) effect on labor productivity growth. Then it decreases and becomes negative (-0.083) and significant (at 10%) two years after the shock, and it continues to be negative (-0.141, -0.188, -0.228) and significant (at 1%) three, four, and five years after the shock.

7.2 Cross-sectional IV regressions

Although in the previous subsection analysis, the data and model show similar correlations between labor force growth and labor productivity growth, this analysis does not allow us to determine the causal impact of labor force growth on labor productivity growth. For example, one possible explanation for our result is that workers relocate to states with greater expected labor productivity growth. Nonetheless, our mechanism for the effect of labor force growth on labor productivity growth is quite specific, as it is based on a decrease in the number of new businesses. Karahan, Pugsley and Sahin (2019) validates this mechanism by demonstrating a causal relationship between labor force growth and the number of startups or the startup rate. In particular, they identify that a 1-percentage-point decrease in the working-age population growth rate roughly translates to a nearly 1-percentage-point decrease in the startup rate. Because our model was constructed using a firm-dynamics model similar to Karahan, Pugsley and Sahin (2019)’s framework, it is not surprising that we attain the same kind of relationship, as shown in Figure 6 panels (A) and (D).

Karahan, Pugsley and Sahin (2019), inspired by Shimer (2001), used a past birth rate as an instrumental variable for labor force growth. This variable is a powerful instrument because, as previous research has shown, there is a close connection between current labor force growth and the birth rate some 20 years ago. Furthermore, in our scenario, the birth rate many years ago is unlikely to directly impact current labor productivity growth. Unfortunately, we find that this instrument is too weak to be used for the yearly dynamics across states examined above using
local projections. Lagged state-level birth rate, on the other hand, is a significant predictor of differences in labor force growth after averaging state data over a 30-year span. This fact enables us to use cross-sectional regressions in an attempt to identify the causal effect of labor force growth on labor productivity growth.

For the 30 years from 1990 to 2019, we average labor productivity growth, $g_{prod}$, and labor force growth, $g_M$. We use the birth rate pushed back 15 years as an instrument for $g_M$, so the average is from 1975 to 2004.32 We control by two potentially important variables. First, we control by the state's initial income per capita (average from 1986 to 1989) because state convergence would suggest a negative link between the initial level of development and future growth. Second, we include the state’s population (average from 1990 to 2019), as many growth theories indicate that scale effects may exist.

The findings of six specifications are presented in Table 5. The first three columns show the results of OLS regressions, while the next three columns show the results of Instrumental Variable (IV) regressions. The three alternative specifications for each approach differ in how state-level observations are weighted. In the first instance, all observations are weighted equally. The weight in the second case is the logarithm of the state’s population, and the weight in the third case is the state’s population.

Table 5, regardless of specification, demonstrates that labor force growth has an effect on labor productivity of around 0.2 percentage point change in labor productivity growth for every 1-percentage-point change in labor force growth. It should be noted that these estimates are comparable to the effect estimated in the local projections for the years following the shock. In general, the estimates in the OLS regressions are more significant than in the IV regressions, though all of the coefficients for the effect of $g_M$ on $g_{prod}$ are significant at a 5% level. Furthermore, except when states are weighted by population, the F statistics of the first-stage regres-

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32 We also perform the analysis using the average between 1975 and 1989 to avoid period overlap for the averages, and the results are very similar.

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Table 5: Impact of labor force growth on labor productivity growth, cross-sectional regressions for US states

<table>
<thead>
<tr>
<th>Dependent variable</th>
<th>OLS state’s weight</th>
<th>IV, lagged birth rate state’s weight</th>
<th>Population</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Average $g_{prod}$</td>
<td>Equal log(Population)</td>
<td>0.202 (0.044)</td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.000)</td>
<td>(0.000)</td>
</tr>
<tr>
<td></td>
<td>log(Initial income pc)</td>
<td>0.007 (0.026)</td>
<td>0.007 (0.026)</td>
</tr>
<tr>
<td></td>
<td>log(Population)</td>
<td>-0.001 (0.239)</td>
<td>-0.001 (0.276)</td>
</tr>
<tr>
<td></td>
<td>R-squared</td>
<td>0.310 (0.290)</td>
<td>0.312 (0.267)</td>
</tr>
<tr>
<td></td>
<td>First-stage reg F stat</td>
<td>–</td>
<td>–</td>
</tr>
<tr>
<td></td>
<td>Observations</td>
<td>49</td>
<td>49</td>
</tr>
</tbody>
</table>

Note: There is also a constant in each regression, and the values in parenthesis are the p-values corresponding to robust standard errors. The states include all US states except Alaska, Hawaii, and the District of Columbia.

Estimations show that the lagged birth rate is a very powerful instrument. Note also that the magnitudes are comparable, although slightly smaller, than those estimated by Peters (2022) using forced population expulsions in post-war Germany.

### 7.3 The decline in US business dynamism

In this section, we examine how this model predicts the decline in business dynamism in the United States. Since we do not use trend data for the variables we compare, it may be challenging for the model to reproduce the US decline in business dynamism once we feed it the labor force trend. In that sense, the exercise in this section serves as a model validation exercise.

We use the calibration in the previous section to compute the model’s transitional dynamics. To begin, we compute a new BGP based on 1970 population growth. Then, we compute a transition in which the agents are aware of the change in $g_M$ exactly 100 years in advance. Thus, agents can predict the change—this appears appealing because low-frequency population size changes are predictable. Since we have $g_M$ data until 2020, we assume that after that, $g_M$ remains constant at the 2020 level.
Because the time series for these variables in the US begins around 1980, the three plots in Figure 8 begin with the model’s prediction in 1970. Overall, Figure 8 shows that the model, at least to some extent, captures the decline in business dynamism in the US. The model predicts the decline in the exit rate, the entry rate, and the proportion of young establishments. It also predicts a slight increase in the average business size. Note that it is reassuring that the model performs reasonably well in this dimension. The mechanism is straightforward, and it is similar to that described in Karahan, Pugsley and Sahin (2019) and Hopenhayn, Neira and Singhania (2022). Because fewer establishments enter the market (bottom left panel) as population growth slows, the share of young establishments (bottom right panel) decreases. Then, since older establishments are less likely to exit and larger, the exit rate falls (top right panel), and the average business size increases (top left panel). Previous research has shown a link between population growth and business dynamism, so our interest in business dynamism is only a validation exercise.

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33We define young establishments as age 5 or younger.
8 Discussion: Our mechanism vis-à-vis scale effects

We have abstracted away from scale effects on growth up to this point by extending the firm-dynamics model of Hopenhayn (1992). However, those models have a long history, as elegantly explained by Jones (2022). In this section, we briefly discuss a particular kind of model with scale effects, a model in which each business produces a different type of good or variety.

Assume the technology for a project of productivity \( x_i \), capital \( k_i \) and labor \( l_i \) is

\[ y_i = x_i k_i^{\alpha_i} l_i^{1-\alpha_i} \]

and the final consumption good is a CES combination of goods or varieties according to

\[ Y = \left( \sum_{i=1}^{N} y_i^{\sigma} \right)^{\frac{1}{\sigma-1}}, \]

where \( N \) is the number of firms, each producing a different variety as in Peters and Walsh (2021). These expressions give the following formula for total output,

\[ Y = N^{\frac{1}{(\sigma-1)}} \tilde{X} K^\alpha M^{1-\alpha}, \]

where \( \tilde{X} \equiv \left( \sum_{i=1}^{N} x_i^{\sigma-1} \right)^{\frac{1}{\sigma-1}} \) is the CES aggregation of productivity and \( N \) is the number of goods or varieties \((K \text{ and } M \text{ are total capital and labor force, as before})\). Therefore, the growth in TFP is

\[ g_{TFP} = g_{\tilde{X}} + \frac{1}{\sigma-1} g_N. \]

This equation implies that TFP growth is proportional to the growth of average productivity across firms and the growth in the number of businesses/varieties \( g_N \). Recall that for comparisons of balanced-growth paths, in our benchmark model, \( g_{TFP} \) is proportional to \( g_X \), so the last term is the key difference when considering new varieties.

In a BGP, the growth rate in the number of varieties, \( g_N \), must equal the growth rate of the population, \( g_M \). Thus, this equation implies that, unlike in our model, in this model, there is a direct effect of population growth on TFP growth across
BGPs. The magnitude of this effect is given by the parameter $\sigma$. Calibrating $\sigma = 4$ such that it is consistent with the “degree of diminishing returns” calibrated in Jones (2022), this equation says that for each 1-percentage-point decline in population growth, there would be a 0.33 percentage point decline in productivity growth. This number is larger but comparable with the numbers we found for the change in $g_X$: 0.13 for the US’s calibration and 0.21 for Japan’s calibration. Thus, the new mechanism introduced in this paper would increase the impact of population growth on productivity growth between 40-60% (from 0.36 to 0.42-0.54 percentage points). Similarly, Peters and Walsh (2021) finds that in the long run, for each point of decline in population growth, productivity growth declines by about 0.23 percentage points. They also report that almost all of that is because of the decline in the number of varieties, which is the term highlighted in this section. Thus, the impact we find for our new mechanism would increase their effect on productivity growth in the US by 56% (from 0.23 to 0.36 percentage points).

9 Conclusions

At least since Solow (1957), the persistent improvement in living standards around the globe has been largely attributed to TFP growth. However, the trend growth in TFP has recently slowed in industrialized economies (Cette, Fernald and Mojon, 2016; Fernald et al., 2017). On the other hand, population growth has declined in most developed economies, and this trend is projected to continue in the next decades. In fact, the latest United Nations projections suggest that the world’s population could reach zero growth during the 2080s (UN, 2022). Therefore, figuring out the possible impact of population growth on TFP growth is crucial. We offer a theory that ties these two trends together. According to our theory, the slowdown in population growth has been, and will likely continue to be, a drag on TFP growth in the coming decades due to the projected slowing population growth.
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A  Appendix

A.1  Proofs

While we assume in section 2 that the success probability $\lambda$ and survival rate of unsuccessful businesses $s_U$ are constant, all the lemmas are proven under a more generalized case here. Specifically, we assume that they are weakly decreasing in age, which is consistent with the calibration in section 4.2.

A.1.1  Proof of Lemma 1

First, the value of a project with potential productivity $\hat{x}_t$ is

$$I_t(\hat{x}_t; \{w_t\}, \{r_t\}) = \sum_{j=t+1}^{\infty} \hat{\beta}_j \mathbb{E}[S(x_j; w_j, r_j) | \hat{x}_t].$$

Note that using equation (2), we have that

$$S_t = \zeta x_t \left(\frac{\alpha}{r_t} \left(1 - \alpha - \zeta \right) \left(\frac{1 - \alpha - \zeta}{w_t} \right)\right)^{\frac{1}{\zeta}},$$

so we can rewrite $I_t$ as

$$I_t(\hat{x}_t; \{w_t\}, \{r_t\}) = \sum_{j=t+1}^{\infty} \hat{\beta}_j \mathbb{E}[x_j | \hat{x}_t] = \Gamma(w_t, r_t) \hat{x}_t,$$

where

$$\Gamma(w_t, r_t) = \zeta x_t \left(1 - \frac{1}{\alpha} \frac{1}{\zeta} \right) \left(\frac{1}{w_t} \right)^{\frac{1}{\zeta}} \sum_{j=t+1}^{\infty} \hat{\beta}_j (\Lambda_{S_j}(g_S) + \Lambda_{U_{j-t}}(r_j) \zeta^{-\frac{1}{\zeta}} (w_j)^{-\frac{1}{\zeta}} \right).$$

using equation (6) to substitute for $E[x_j | \hat{x}_t]$.

Then, we can solve equation (3). This equation is altered to

$$V(\{w_t\}, \{r_t\}, \chi_t) = \max_{\sigma_t, \chi_t} \sigma_t G(w_t, r_t) \chi_t - \frac{1}{z_R} (g_t)^t w_t - \frac{\sigma_t^2}{2z_D} w_t,$$ (24)

where $g_t \equiv \hat{x}_t / \chi_t$ is the step size of innovation. Note that the FOCs with respect to
\( \sigma_t \) and \( g_t \) are
\[
\frac{\partial V_t}{\partial \sigma_t} = \Gamma(w_t, r_t)g_t\chi_t - \frac{w_t}{z_D} \sigma_t = 0, \quad \frac{\partial V_t}{\partial g_t} = \Gamma(w_t, r_t)\sigma_t\chi_t - \frac{rw_t}{z_R} g_t^{t-1} = 0.
\]
The solutions are
\[
g^*_t = \left( \frac{z_R z D}{l} \Gamma(w_t, r_t) \right)^{\frac{1}{2}} \left( \frac{\chi_t}{w_t} \right)^2, \quad (25)
\]
\[
\sigma^*_t = \frac{z_D \Gamma(w_t, r_t)}{w_t} \chi_t g^*_t. \quad (26)
\]
Substituting equations (25) and (26) into (24), we obtain
\[
V_t = l - 2 \left( \frac{z_R z D}{l} \Gamma(w_t, r_t) \right)^{\frac{1}{2}} \left( \frac{\chi_t}{w_t} \right)^2 w_t = \frac{l - 2}{2z_R} (g^*_t)^t w_t. \quad (27)
\]
Finally, we replace equation (27) into the free entry condition (equation (4)),
\[
V_t = \frac{l - 2}{2z_R} (g^*_t)^t w_t = c_E^* w_t,
\]
which determines the step size \( g^* \),
\[
g^* = \left( \frac{2c_E z_R}{l - 2} \right)^{\frac{1}{t}}.
\]
Since this solution is equal to equation (14), this concludes the proof of Lemma 1.

**A.1.2 Proof of Lemma 2**

Let \( \cdot' \) denote values in the next period. Suppose that \( g_M > s_{S,\infty} \) and \( \{\sigma, g, c, k, N\} \) solves the old problem. The existence of a balanced growth path will be shown when \( \{\sigma, g, g_w c, g_w k, g_M N\} \) solves the new one for \( M' = g_M M \).

First, the Euler equation derived from equation (1), \( (g_w)^c = \beta(1 + r - \delta) \), shows \( r \) is constant, and so \( \hat{\beta}_t' = \hat{\beta}_t \).

Second, when we observe \( N' = g_M N \), we will get a certain value of \( g_x \) from equation (17). Note that the right-hand side of equation (17) is a monotonic (increasing) function of \( g_x \) and \( \lim_{g_x \rightarrow 0} \text{RHS} \rightarrow 0 \) and \( \lim_{g_x \rightarrow \infty} \text{RHS} \rightarrow \infty \) for any \( g_M \in (0, \infty) \). Given the LHS is positive, we have a unique \( g_x \) for every \( g_M \).

Third, the low of motion of the number of businesses, equations (8)-(10), shows
\[ n' = g_{NN}n. \]

In addition, equations (15) and (16) give \( \chi' = g_X \chi \) and \( \hat{x}' = g_X \hat{x} \). Equation (25) gives \( w' = (g_X)^{1-\tau}w \), and it gives \( L' = g_M L \), from the equation for labor demand.

To see the firm side, from equation (2), \( S(\hat{x}', w', r') = g_w S(\hat{x}, w, r) \), and so \( I(\hat{x}', w', r') = g_w I(\hat{x}, w, r) \). Therefore, the objective function for the value of projects in equation (3) inflates by the factor \( g_w \). Note that \( R(\hat{x}' / \chi') = R(\hat{x} / \chi) \) and \( D(\sigma') = D(\sigma) \). Also, both sides of equation (4) inflates by \( g_w \), so the new solution still satisfies the free entry condition.

Regarding market clearing conditions, all labor, capital, and goods markets hold in the new problem at the conjectured solution: both sides of equation (11) inflate by \( g_M \), equation (12) by \( g_w g_M \), and equation (13) by \( g_w g_M \).

Lastly, consider the budget constraint in equation (1). At the conjectured solution, both sides inflate by the factor \( g_w \). As such, the budget constraint also holds with the new solution.

Thus, by canceling out this factor of proportionality, the new problem reverts back to the old one, and we have shown that it is a BGP.

Finally, we show that population growth must be higher than the old businesses’ survival rate, \( g_M > s_{S,\infty} \) to have a BGP. For this, note that the lower bound for the growth in the number of businesses is \( s_{S,\infty} \). This is the case because (i) in that lower bound, the number of entrants is at its lower bound (zero), and (ii) we assumed that successful firms live longer than unsuccessful ones, \( s_{S,\infty} \geq s_{U,\infty} \). Hence, if \( g_M < s_{S,\infty} \), then the population growth rate is lower than the growth rate of the number of businesses, \( g_M < g_N \), which contradicts the condition \( g_M = g_N \) that should hold in any balanced growth path.

A.1.3 Proof of Lemma 3

This proof consists of two parts: (i) Prove that \( g_X > g_S \) if and only if the employment growth rate of surviving successful businesses is negative, and (ii) prove that the employment growth rate of surviving old businesses is asymptotically equiva-
lent to that of surviving *successful* businesses.

First, the employment of a business \( i \) at time \( t \) is

\[
l_{i,t} = x_{i,t} \left[ \left( \frac{\alpha}{r_t} \right)^a \left( \frac{1 - \alpha - \zeta}{w_t} \right)^{1-a} \right]^{\frac{1}{\zeta}}.
\]

Given that the business is successful at time \( t \) and surviving at time \( t + 1 \), the business’s employment at time \( t + 1 \) is written as

\[
l_{i,t+1} = g_S x_{i,t} \left[ \left( \frac{\alpha}{r_{t+1}} \right)^a \left( \frac{1 - \alpha - \zeta}{w_{t+1}} \right)^{1-a} \right]^{\frac{1}{\zeta}}.
\]

Since \( r_{t+1} = r_t \) and \( w_{t+1} = (g_X)^{\frac{\zeta}{1-a}} w_t \) in a BGP, the employment growth surviving of successful businesses \( l_{i,t+1}/l_{i,t} = g_S/g_X \).

Next, to show the asymptotic equivalence between the employment growth rate of surviving *old* businesses and that of surviving *successful* businesses, recall that we assume that old successful businesses are more likely to survive than old unsuccessful businesses, \( s_{S,\infty} \geq s_{U,\infty} \). It implies that the share of unsuccessful businesses converges to zero, so the employment growth rate of old businesses becomes equivalent to that of surviving businesses. This concludes the proof of Lemma 3.

We can see this step algebraically by considering the expression for the employment growth rate of surviving businesses of all ages:

\[
\text{surviving growth} = \frac{g_S}{g_X} (1 - \Delta_a) + \Delta_a \left( \frac{(1 - \lambda_a) + \frac{\lambda_a}{g}}{g_X} \right),
\]

where \( \Delta_a \) is the employment share of unsuccessful businesses. Note that it depends on the employment share, not the business share. Since successful businesses are larger than unsuccessful businesses, the employment share converges faster than the business share.
A.1.4 Proof of Lemma 4

Combining equations (14) and (17) and arranging it gives an equilibrium expression for the relationship between $g_M$ and $g_X$:

$$
(2c_{Ez_R})^{-\frac{1}{\ell - 2}} = \frac{\sum_{a=1}^{\infty} \left( \frac{1}{g_M} \right)^a \left( \frac{g_S}{g_X} \right)^a \Lambda_{S,a}(g_S) \Lambda_{S,a}(1)}{\sum_{a=1}^{\infty} \left( \frac{1}{g_M} \right)^a \Lambda_{S,a}(1)}.
$$

The right-hand side can be interpreted as the weighted average of $\tilde{x}_a \equiv \left( \frac{g_S}{g_X} \right)^a \Lambda_{S,a}(g_S) / (g_S)^a \Lambda_{S,a}(1)$ with weights $(1/g_M)^{a-1} \Lambda_{S,a}(1)$; i.e.,

$$
(2c_{Ez_R})^{-\frac{1}{\ell - 2}} = \sum_{a=1}^{\infty} (\tilde{x}_a) \times \frac{((1/g_M)^a \Lambda_{S,a}(1))}{\sum_{a=1}^{\infty} (1/g_M)^a \Lambda_{S,a}(1)}.
$$

(28)

Now, $\tilde{x}_a$ is decreasing in $g_X$ for all $a \geq 1$. As such, ceteris paribus, the increase in $g_X$ decreases the right-hand side of equation (28). Also, if $g_S < g_X$, $\tilde{x}_a$ is decreasing in age $a$ since $\Lambda_{S,a+1}(g_S) / \Lambda_{S,a}(g_S) \leq g_S \Lambda_{S,a+1}(1) / \Lambda_{S,a}(1)$. Therefore, we need larger weights on young businesses to increase the right-hand side of equation (28) if $g_S < g_X$, which means that we need to increase $g_M$.

To put them together, if $g_S < g_X$, an increase in $g_M$ must increase $g_X$ to keep the right-hand side of equation (28) constant, which concludes the proof of Lemma 4.

A.1.5 Proof of Lemma 5

Start with equation (17). Totally differentiating it by $g_M$ and reorganizing it gives

$$
\frac{1}{1 + \frac{g_M}{g_X} d g_X / d g_M} \sum_{a=1}^{\infty} \left( \frac{1}{g_M} \right)^a \Lambda_{S,a}(1) = \sum_{a=1}^{\infty} \left( \frac{1}{g_M} \right)^a \left( \frac{1}{g_X} \right)^a \Lambda_{S,a}(g_S) a.
$$

(29)

Therefore, for given $g_M$ and $g_X$, $d g_X / d g_M$ is larger if the right-hand side of equation (29) is smaller. Note that only the right-hand side depends on $g_S$. Since the growth rate in the size of old businesses decreases faster in the economy with
smaller $g_s$, we need to show
\[
\frac{d}{dg_s} \sum_{a=1}^{\infty} \left( \frac{1}{S_a} \right)^a \left( \frac{1}{S_x} \right)^a \Lambda_{S,a}(g_s) a > 0.
\]

Arranging it gives
\[
\frac{\sum_{a=1}^{\infty} \left( \frac{1}{S_a} \right)^a \left( \frac{1}{S_x} \right)^a \Lambda_{S,a}(g_s) a \left( \frac{\Lambda'_{S,a}(g_s)}{\Lambda_{S,a}(g_s)} \right)}{\sum_{a=1}^{\infty} \left( \frac{1}{S_a} \right)^a \left( \frac{1}{S_x} \right)^a \Lambda_{S,a}(g_s)} > \frac{\sum_{a=1}^{\infty} \left( \frac{1}{S_a} \right)^a \left( \frac{1}{S_x} \right)^a \Lambda_{S,a}(g_s) \left( \frac{\Lambda'_{S,a}(g_s)}{\Lambda_{S,a}(g_s)} \right)}{\sum_{a=1}^{\infty} \left( \frac{1}{S_a} \right)^a \left( \frac{1}{S_x} \right)^a \Lambda_{S,a}(g_s)},
\]
where $\Lambda'_{S,a}(g_s) \equiv d\Lambda_{S,a}(g_s)/dg_s$. Note that both sides of equation (30) can be interpreted as weighted averages of $\Lambda'_{S,a}(g_s)/\Lambda_{S,a}(g_s)$. Since LHS of equation (30) has more weights on older ages, a sufficient condition for this equation to hold is $\Lambda'_{S,a}(g_s)/\Lambda_{S,a}(g_s)$ is increasing in age $a$.

Now, using the definition of $\Lambda_{S,a}(g_s)$, we can write $\Lambda'_{S,a}(g_s)/\Lambda_{S,a}(g_s)$ as
\[
\frac{\Lambda'_{S,a}(g_s)}{\Lambda_{S,a}(g_s)} = \frac{\sum_{j=1}^{a} (j-1) \omega_{a,j}}{\sum_{j=1}^{a} \omega_{a,j}},
\]
where $\omega_{a,j} \equiv g s^{j-1} \left( \prod_{k=0}^{a-j-1} (1 - \lambda_k) \right) \left( \prod_{k=1}^{a-j} s_{U,k} \right) \left( \prod_{k=a-j+1}^{a-1} s_{S,k} \right) \lambda_{a-j}$. Notice that the subscript $j$ represents how many periods have passed since each business becomes successful, so $\omega_{a,j}$ is the aggregate productivity of successful businesses at age $a$ that succeeded at age $a - j + 1$. Therefore, what we want to show gets to
\[
\sum_{j=1}^{a+1} (j-1) \omega_{a+1,j} > \sum_{j=1}^{a} (j-1) \omega_{a,j}
\]
for all ages $a$. Using a property $\omega_{a+1,j+1} = g s^s S_{S,a} \omega_{a,j}$, we can transform equation (31) into
\[
\sum_{j=1}^{a} \omega_{a,j} > \sum_{j=1}^{a} \lambda_a
\]
Since we assume that the survival probability for successful businesses is higher than that for unsuccessful $s_{S,a} > s_{U,a}$ for all ages and the successful probability $\lambda_a$
decreases in age \( a \), \( \omega_{a,j+1} > \omega_{a,j} \) because
\[
\frac{\omega_{a,j+1}}{\omega_{a,j}} = \frac{g_s s s_{a-j} \lambda_{a-j-1}}{(1 - \lambda_{a-j-1}) s U_{a-j} \lambda_{a-j}}.
\]

Under these assumptions, we can prove equation (32):
\[
(LHS) = \frac{\sum_{j=1}^{a} \omega_{a,j}}{\omega_{a+1,1} / g_s s S_s a} > \frac{\sum_{j=1}^{a} \omega_{a,j}}{\omega_{a,1}} > \frac{\sum_{j=1}^{a} \omega_{a,j}}{\omega_{a,j}/a} = a > \frac{\sum_{j=1}^{a} (j - 1) \omega_{a,j}}{\sum_{j=1}^{a} \omega_{a,j}} = (RHS).
\]

Hence, \( dg X/dg M \) is larger in the economy in which the growth rate in the size of old businesses decreases faster.

### A.2 Effect of population growth on the share of capital \( \tilde{\alpha} \)

Starting since the definition of \( \tilde{\alpha} \), note that since \( wL/Y = 1 - \alpha - \zeta \) from the definition of technology, we an obtain
\[
\tilde{\alpha} = 1 - (1 - \alpha - \zeta) \frac{M}{L},
\]

On the business side, the following equation holds:
\[
Y = rK + wL + S, \quad (33)
\]

Note that \( rK = \alpha Y \), \( wL = (1 - \alpha - \zeta) Y \), and \( S = \zeta Y \).

Conversely, on the household side, we have
\[
Y = rK + wM + S - E. \quad (34)
\]

Note that \( w(M - L) = E \) because entry, research, and development require only labor.

Combining equations (33) and (34), we derive
\[
\frac{M}{L} = 1 + \frac{\zeta}{1 - \alpha - \zeta} \frac{E}{S},
\]

so the ratio of \( E \) to \( S \) affects the share of workers \( \tilde{\alpha} \) through the ratio of \( M \) to \( L \).

Additionally, in the context of the free entry condition, the initial cost of starting a business \( E \) equals the expected future profit that entrants will earn until they exit. Therefore, the ratio \( E/S \) represents the ratio of expected future profit to current
profit. Along a BGP, this ratio transforms into:

\[
\frac{E}{S} = \frac{\sum_{t=1}^{\infty} \left( \frac{g \omega}{1 + r - \delta g_X} \right)^t (\Lambda_{S,t}(g_S) + \Lambda_{U,t})}{\sum_{t=1}^{\infty} \left( \frac{1}{g M g_X} \right)^t (\Lambda_{S,t}(g_S) + \Lambda_{U,t})}.
\]

Given that \( dg_X/dg_M \) is substantially smaller than one both from model and empirical results, an increase in \( g_M \) decreases the denominator more than the numerator, as long as \( \frac{\xi (e-1)}{1-\alpha} \) is not large. Consequently, \( d\tilde{\alpha}/dg_M \) is likely negative, amplifying the effect of population growth on TFP growth, although extreme parameter values can alter this relationship.

### A.3 Spillovers calibration

In this section, we estimate the value of \( \gamma \) for equation (19). To obtain a measure of \( g_S \), we use BDS data and the following procedure. Use the data to construct the share of old establishments \( (N_{old}/N) \) and the share of workers in old establishments \( (L_{old}/L) \), where we consider an establishment as old if it is 16 years old or older. Then, note that from the equation for aggregate labor in production \( L \), we can construct data on \( \log(X_{old}/X_t) \) as it is equal to \( \log([L_{old}/L]_t) - \log([N_{old}/N]_t) \). Finally, we obtain \( g_S \) as \( g_S = \Delta \log([X_{old}/X]_t) + \Delta \log(X_t) = \Delta \log ([X_{old}/X]_t) + g_X. \)

The OLS estimation of equation (19) is in the first column of Table 6. That is the coefficient we use in our benchmark model. The second column is the same regression but adds a linear trend. It yields similar results.
Table 6: Estimates for calibration of spillovers

<table>
<thead>
<tr>
<th>Regression for $g_s$</th>
<th>OLS</th>
<th>Instrumental Variables</th>
</tr>
</thead>
<tbody>
<tr>
<td>$g_{x,t-1}$</td>
<td>0.342*</td>
<td>0.384* 0.387 0.458** 0.466**</td>
</tr>
<tr>
<td></td>
<td>(0.186)</td>
<td>(0.207) (0.205) (0.196)</td>
</tr>
<tr>
<td>Trend</td>
<td>no</td>
<td>yes no yes no yes</td>
</tr>
<tr>
<td>R squared</td>
<td>0.124</td>
<td>0.108 0.092 0.115 0.093</td>
</tr>
<tr>
<td>First stage statistic F</td>
<td>-</td>
<td>14.380 11.930 25.190 21.540</td>
</tr>
<tr>
<td>Hansen’s $\chi^2$, p value</td>
<td>-</td>
<td>0.128 0.144 0.168 0.194</td>
</tr>
<tr>
<td>Instruments</td>
<td>-</td>
<td>VC VC &amp; entry VC &amp; entry</td>
</tr>
<tr>
<td>Observations</td>
<td>23</td>
<td>23 23 23 23</td>
</tr>
</tbody>
</table>

Note: “VC” stands for lag growth rate of (i) VC total investment, (ii) early stage investment, (iii) seed investment, and (iv) expansion stage investment. Similarly, “Entry” stands for lag growth rate in the entry rate.

One may be concerned that overall productivity growth in $t - 1$ may be affected by the productivity of the old businesses in period $t$. Ideally, we want the variation in $g_X$ that is independent of the productivity growth of already successful businesses. We chose it as an instrument of $g_X$ venture capital (VC) investment because it should affect $g_X$ through innovation by new firms. And VC investment should not directly affect the productivity growth of already successful businesses. Using this IV approach, we find slightly larger estimates.

Finally, it may be that old businesses’ productivity growth causes VC investment. For example, if entrants learn from incumbents and VCs respond to the amount of spillover from incumbents to entrants. This relationship will break, however, if we consider VC investment that is lagged with respect to aggregate productivity growth. Although is not reported, we also consider a first stage in which 3 to 8 lags of seed-stage VC investment are the instruments for aggregate productivity growth. The first stage F is 13.34, and our coefficient of interest is 0.502, close to the estimates in the last columns of Table 6.
A.4 Profiles of survival and success probabilities

Figure 9: Probability of survival and success over the life-cycle
B Online Appendix

B.1 Data sources

B.1.1 US data


Establishment data. Establishment data come from the U.S. Census Bureau’s Business Dynamics Statistics (BDS). It provides annual measures of establishment openings and closings, and job creation and destruction by age group, which allows computing life-cycle patterns of establishment that we use as targets in Figure 3 and other dynamism in Figure 8. The data are available since 1978. An establishment is a fixed physical location where economic activity occurs. A firm may have one establishment or many establishments.

Total factor productivity. Total factor productivity (TFP) data are calculated using Penn World Table (PWT) 10.0. While the value of TFP is directly available in PWT, we compute TFP from real GDP, number of persons engaged, and capital stock by assuming a Cobb-Douglas function. It ignores the effect of the change in human capital, which makes the data more consistent with our model. The data have been available since 1950.

Venture capital investment. Venture capital investment data are from the PwC/CB Insights MoneyTree™ Report.
B.1.2 Japan data

**Labor force.** Labor force data come from the Statistics Bureau of Japan (SBJ) Labor Force Survey from 1953 to 2019 and the National Institute of Population and Social Security Research (IPSS) from 1920 to 1952. The definition includes the population 15 years of age and over. The labor force growth projections are taken from the Ministry of Internal Affairs and Communications (MIC) "Information and Communications in Japan 2022," published in July 2022. It offers only a projection of the working-age population, so we assume the ratio of the labor force to the working-age population has remained constant since 2020.

**Establishment data.** Establishment data come from SBJ’s Establishment and Enterprise Census from 1981 to 2006 and from Economics Census from 2009 to 2021. They provide the number of establishments and employment by age group while they are not annual data (available only in 1981, 1986, 1991, 1996, 2001, 2004, 2006, 2009, 2012, 2014, 2016, and 2021). The number of establishments and employment by both age groups and three kinds of status (existing, newly organized, and closed) are only available in 2004, so the exit rate and growth of surviving establishments by age group, which are targets, are calculated based on the data in 2004 by comparing with the data in 2001. Accurately, the calculation will not give the annual exit rate and growth but a three-year average. Therefore, the fittings for these targets in Figure 3 are based on this three-year average at an annual rate.

We have extracted the birth-year-fixed effect for employment size by age, another target value. The life-cycle profile for Japan is largely influenced by the birth year, while it remains very stable over time in the United States. In particular, we first assume that all ages in the same age group grow their establishment size at the same rate over the year. (For example, the establishment size for the age group between 3 and 7 grew by 1.2% annually from 2001 to 2004; we assume age 3, 4, 5, 6, and 7 establishments in 2001 increased their size by 1.2% every year between 2001 and 2004.) Then, we regress establishment size growth on fixed effects of age $a$ and...
year $t$: Establishment size growth $a,t = u_a + v_t + \epsilon_{a,t}$, and extract $u_a$ to get average establishment size growth by age. Please note that with age $a$ and year $t$, the born year is identically defined by $t - a + 1$.

**Total factor productivity.** TFP data are calculated using PWT 10.0, as in the US case.

### B.1.3 US data (state-level)

**Civilian labor force.** Civilian labor force data come from the BLS Current Population Survey for the years 1976 to 2019 and from *Historical Statistics of the United States, Colonial Times to 1970* issued by the Census Bureau for the years 1900 to 1950. We interpolate the data between 1950 and 1976. The projections for 2030 are taken from the Projections Managing Partnership (PMP), which is based on employment. We use only data between 1976 and 2019 to estimate empirical data, where yearly data are available, while we use the full data to estimate simulated data to reproduce TFP growth.

**Real GDP and total nonfarm employees.** Real GDP data come from the Bureau of Economic Analysis (BEA) Gross Domestic Product by State and Personal Income by State. Total nonfarm employees for the corresponding years are from the BLS. We use these statistics to derive labor productivity.

**Population.** Population data are from the Census Bureau. This is one of the control variables in the cross-sectional regressions for US states.

**Real income per capita.** Real income per capita data, another control variable for the cross-sectional regressions, come from the BEA Real Personal Income for States and Metropolitan Areas between 2008 and 2019, and from the Census Bureau’s Statistical Abstract of the United States from 1975 to 2007.
B.2 More on BGP for the benchmark case

Section 3 shows the balanced growth path without two extensions and adds two features (congestion and spillover) thereafter. However, these two features somewhat change the property of the balanced growth path.

First, since the productivity growth of successful projects $g_S$ depends on $g_X$, which is equal to $g_X$ along a balanced growth path, due to spillover, expected productivity for successful projects $\Lambda_S(g_S)$ will depend on $g_X$. Let it denote as $\Lambda_S(g_S(g_X))$. As such, the RHS of the equation for the step size $g^*$ (17) is written in this form:

$$g^* = \frac{\sum_{a=1}^{\infty} \left( \frac{1}{SM} \right)^a \Lambda_S(1)}{\sum_{a=1}^{\infty} \left( \frac{1}{SXM} \right)^a \Lambda_S(g_S(g_X))}. \quad (35)$$

Next, because of congestion, the LHS of equation (17) (or equation (14)) is altered to

$$g^* = \left( \frac{2c_E(\tilde{n}/\tilde{M})^\phi z_R^t}{1 - 2} \right)^{\frac{1}{7}}, \quad (36)$$

so we need to specify $n/M$. Using the equation for labor demand, we can derive

$$\frac{N}{L} = \frac{w^{1-\alpha}}{X} \left( \frac{\alpha}{r} \right)^{-\frac{2}{7}} (1 - \alpha - \zeta) \frac{1-\alpha}{7}. \quad (37)$$

Also, from equation (25),

$$w^{\frac{1-\alpha}{7}} = \chi(g^*)^{-\frac{1-\alpha}{7}} \left( \frac{ZRZD}{t} \right)^{\frac{1}{7}} \zeta \left( \frac{\alpha}{r} \right)^{\frac{2}{7}} (1 - \alpha - \zeta) \frac{1-\alpha}{7} \times$$

$$\sum_{t=1}^{\infty} \hat{\beta}_t(\Lambda_S(g_S) + \Lambda_U(t)) \left( \frac{1}{g^w} \right)^{\frac{1-\alpha-\zeta}{7}(t-1)}. \quad (38)$$

Note that

$$\Gamma(w, r) = \zeta \left( \frac{\alpha}{r} \right)^{\frac{1}{7}} \left( \frac{1 - \alpha - \zeta}{w} \right)^{\frac{1-\alpha-\zeta}{7}} \sum_{t=1}^{\infty} \hat{\beta}_t(\Lambda_S(g_S) + \Lambda_U(t)) \left( \frac{1}{g^w} \right)^{\frac{1-\alpha-\zeta}{7}t}$$
along a balanced growth path. Therefore, from equations (37) and (38),

$$\frac{N}{L} = \frac{X}{X}(g^*)^{-\frac{1-a}{\alpha}} \left( \frac{ZRD}{t} \right)^{\frac{1}{2}} \left( 1 - \alpha - \zeta \right) \sum_{t=1}^{\infty} \hat{\beta}_t (\Lambda_{S,t}(g_S) + \Lambda_{U,t}) \left( \frac{1}{g_w} \right)^{\frac{1-a-\zeta}{\zeta}}. \quad (39)$$

Since

$$N = n \left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a (\Lambda_{S,a}(1) + \Lambda_{U,a}/\theta) \right)$$

from equations (8) and (9), and

$$\frac{X}{X} = \frac{\left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(1) \right) \left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(g_S) \right)}{\left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(1) \right) \left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(g_S) + \Lambda_{U,a} \right)}$$

from equations (15) and (16), equation (39) gives this equation:

$$\frac{n}{L} = (g^*)^{-\frac{1-a}{\alpha}} \left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(g_S) \right) \left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(1) \right) \times$$

$$\left[ \left( \frac{ZRD}{t} \right)^{\frac{1}{2}} \left( 1 - \alpha - \zeta \right) \left( \sum_{t=1}^{\infty} \left( \frac{\beta}{g_w} \right)^t \left( \Lambda_{S,t}(g_S) + \Lambda_{U,t} \right) \right) \right].$$

Substituting $g^*$ for the RHS of equation (17) and $g_w$ for $(g_X)^{\frac{1}{1-\alpha}}$, we can derive the following expression for $n/L$:

$$\frac{n}{L} = \left( \frac{ZRD}{t} \right)^{\frac{1}{2}} \left( 1 - \alpha - \zeta \right) \times \left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(g_S) \right) \left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(1) \right) \times$$

$$\left[ \left( \frac{\beta}{g_X} \right)^{t} \left( \Lambda_{S,t}(g_S) + \Lambda_{U,t} \right) \right]. \quad (40)$$

$$\left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(g_S) \right) \left( \sum_{a=1}^{\infty} \left( \frac{1}{S_M} \right)^a \Lambda_{S,a}(1) \right) \times$$

$$\left[ \left( \frac{\beta}{g_X} \right)^{t} \left( \Lambda_{S,t}(g_S) + \Lambda_{U,t} \right) \right]. \quad (41)$$
Also, from equations (4), (11), (25), (26), and (27),

\[
\frac{M}{L} = 1 + \frac{\zeta}{1 - \alpha - \zeta} \left( \sum_{t=1}^{\infty} \left( \frac{\beta}{g_X^{1/(1-\alpha)}} \right)^t \left( \Lambda_{S,t}(g_S(g_X)) + \Lambda_{U,t} \right) \right) \left( \sum_{a=1}^{\infty} \left( \frac{1}{s_Xs_M} \right)^a \left( \Lambda_{S,a}(g_S(g_X)) + \Lambda_{U,a} \right) \right)^{-1}. \tag{42}
\]

Therefore, from equations (41) and (42), the step size \( g^* \) is influenced by \( g_M \) and \( g_X \) through the number of entrants per capita. These equations and the equation for the step size

\[
\left( \frac{2c_E(n/M)\psi_z}{t-2} \right)^{\frac{1}{t}} = \frac{\sum_{a=2}^{\infty} \left( \frac{1}{s_M} \right)^{a-1} \Lambda_{S,a}(1)}{\sum_{a=2}^{\infty} \left( \frac{1}{s_Xs_M} \right)^{a-1} \Lambda_{S,a}(g_S(g_X))}
\]

define the equilibrium \( g_X \) for each \( g_M \).

**B.3 Sensitivity analysis**

We simulate a shock of \( g_M \) from 1.02 to 1.01 to \( g_{TFP} \) using the US calibration as a benchmark case. The impact size is the change in \( g_{TFP} \), in the long run, resulting from that change. The elasticity of "Impact Size" represents the elasticity of the impact of \( g_M \) on \( g_{TFP} \) to a change in a parameter value. Specifically, an "Impact Size" elasticity equal to \( X \) means that a 1% change in the parameter results in an \( X\% \) increase in the response of \( g_M \) to \( g_{TFP} \).

"Convergence Speed" is the share of the change in \( g_{TFP} \) that occurred 20 periods after the \( g_M \) shock (relative to the long-run impact). The “Convergence Speed” elasticity equal to \( Z \) means that a 1% change in the parameter results in a \( Z\% \) increase in the explained share of \( g_{TFP} \) change 20 periods after the shock.

Finally, note that for the parameters that vary over age, the exact change to \( \lambda \) and \( s_U \) is applied to all ages.
Table 7: Role of parameters for the impact size and the convergence speed

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Elasticity</th>
<th>Impact Size</th>
<th>Convergence Speed</th>
</tr>
</thead>
<tbody>
<tr>
<td>Survival rate of successful businesses, $s_S$</td>
<td></td>
<td>12.819</td>
<td>-2.479</td>
</tr>
<tr>
<td>Survival rate of unsuccessful businesses, $s_U$</td>
<td></td>
<td>2.384</td>
<td>-3.180</td>
</tr>
<tr>
<td>Decreasing returns, $\zeta$</td>
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<td>1.039</td>
<td>1.142</td>
</tr>
<tr>
<td>Capital share, $\alpha$</td>
<td></td>
<td>0.571</td>
<td>0.972</td>
</tr>
<tr>
<td>Spillover elasticity, $\gamma$</td>
<td></td>
<td>0.244</td>
<td>0.293</td>
</tr>
<tr>
<td>Entry cost exponent, $\phi$</td>
<td></td>
<td>0.184</td>
<td>0.129</td>
</tr>
<tr>
<td>Success probabilities, $\lambda$</td>
<td></td>
<td>0.043</td>
<td>0.083</td>
</tr>
<tr>
<td>Depreciation rate, $\delta$</td>
<td></td>
<td>0.016</td>
<td>0.276</td>
</tr>
<tr>
<td>Risk aversion, $\epsilon$</td>
<td></td>
<td>0.000</td>
<td>-0.160</td>
</tr>
<tr>
<td>Inverse of Jump in prod at success, $\theta$</td>
<td></td>
<td>-0.052</td>
<td>-0.086</td>
</tr>
<tr>
<td>Research cost exponent, $\iota$</td>
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<td>-0.182</td>
<td>0.128</td>
</tr>
<tr>
<td>Discount factor, $\beta$</td>
<td></td>
<td>-0.241</td>
<td>-0.364</td>
</tr>
<tr>
<td>Productivity growth old businesses, $g_S$</td>
<td></td>
<td>-31.031</td>
<td>-26.141</td>
</tr>
</tbody>
</table>

B.4 Computational Details

In this section, we describe how we compute the model.

B.4.1 Balanced Growth Path

We first solve for a BGP using the 1980-1999 population and productivity growth averages as targets. Substituting these target values for $g_M$ and $g_X$ in equation (35) gives the step size $g^*$ for the reference periods. We then derive the initial productivity growth to define the BGP state before the population growth shock. Given the step size and the discussion in the online appendix B.2, we can find an equation for $g_X$ as a function of the exogenous value $g_M$. Since an analytical solution does not exist, we use Newton’s method. We can also derive the other variables using the properties of the BGP as described in Lemma 2.
B.4.2 Transitional Periods

After computing the BGP, we compute the transitional dynamics in two steps. First, we guess the growth rate of the number of entrants \( \{g_n_t\} \) and solve for the equilibrium set of prices \( \{g_w_t, r_t\} \) using the capital, goods, and labor market clearing conditions. We then derive \( \{g_n_t\} \) using the free entry condition. Each step requires finding the roots numerically, so our code has a nested structure of two Newton’s method computations. Importantly, we solve all periods simultaneously, not sequentially; innovators stand on the shoulders of previous innovators, which requires solving the model from the past to the future. However, they must also choose the step size of productivity, taking expected profits into consideration, which requires solving the model from the future to the past.

In the first step, we derive the number of businesses by age given the conjectured \( \{g_n\} \) and the values in the initial BGP. We can also determine the productivity by age since the step size satisfies equation (36) even during transitional periods. Next, we guess the interest rate \( \{r_t\} \). As the number of businesses \( N_t \) and average productivity \( X_t \) are already known, \( \{g_w_t\} \) can be derived easily by the labor market clearing condition (11) given \( \{r_t\} \). With this set of prices \( \{g_w_t, r_t\} \), we solve the household and business optimization problems. Lastly, we fix \( \{r_t\} \) using Newton’s method using capital and goods market clearing conditions.

In the second step, we compute the expected profits for entrants and, therefore, the value of businesses. Since it should equal the entry cost, we adjust \( \{g_n_t\} \) using Newton’s method. Since changes in \( \{g_n_t\} \) affect \( \{r_t\} \), we return to the first step for every repetition. When the value of businesses minus the entry cost converges to zero sufficiently (\(< 10^{-6}\) ), the computation is finished.

B.5 Full transitions

This section shows the full transitions computed for the US and Japan. As mentioned in the main text, the transition’s input is the labor force growth trend,
which is shown with a black line in Figure 10.

Figure 10: Full Transitions

United States

Japan